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Addressing Statistical Biases in Nucleotide-Derived Protein Databases for Proteogenomic Search Strategies

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Supporting Information

ABSTRACT: Proteogenomics has the potential to advance genome annotation through high quality peptide identifications derived from mass spectrometry experiments, which demonstrate a given gene or isoform is expressed and translated at the protein level. This can advance our understanding of genome function, discovering novel genes and gene structure that have not yet been identified or validated. Because of the high-throughput shotgun nature of most proteomics experiments, it is essential to carefully control for false positives and prevent any potential misannotation. A number of statistical procedures to deal with this are in wide use in proteomics, calculating false discovery rate (FDR) and posterior error probability (PEP) values for groups and individual peptide spectrum matches (PSMs). These methods control for multiple testing and exploit decoy databases to estimate statistical significance. Here, we show that database choice has a major effect on these confidence estimates leading to significant differences in the number of PSMs reported. We note that standard target:decoy approaches using six-frame translations of nucleotide sequences, such as assembled transcriptome data, apparently underestimate the confidence assigned to the PSMs. The source of this error stems from the inflated and unusual nature of the six-frame database, where for every target sequence there exists five “incorrect” targets that are unlikely to code for protein. The attendant FDR and PEP estimates lead to fewer accepted PSMs at fixed thresholds, and we show that this effect is a product of the database and statistical modeling and not the search engine. A variety of approaches to limit database size and remove noncoding target sequences are examined and discussed in terms of the altered statistical estimates generated and PSMs reported. These results are of importance to groups carrying out proteogenomics, aiming to maximize the validation and discovery of gene structure in sequenced genomes, while still controlling for false positives.

KEYWORDS: proteogenomics, peptide spectrum match, false discovery rate, posterior error probability, expressed sequence tag

INTRODUCTION

Rapid advances in mass spectrometry-based proteomics have been made possible by improvements in peptide separation techniques, high-resolution instruments and downstream informatic processing. Researchers can now obtain a comprehensive catalogue of the proteome in single-celled organisms and a near-comprehensive catalogue in multicellular organisms. However, the database search method that underpins the majority of proteomic workflows typically requires a high quality set of protein-coding genes. To improve proteome coverage, mass spectrometry (MS) data can be searched against protein sequences inferred from either the genome or transcriptome. Such proteonomic approaches are not biased toward existing gene annotations and therefore offer scope for novel gene/protein discovery. Indeed, proteogenomics has led to the discovery of thousands of novel gene candidates protein isoforms, amino acid polymorphisms, and confirmation and correction of gene models. A recurring observation in these studies is that current gene models, which are largely computational predictions themselves, are often incomplete and erroneous. For instance, lowly expressed splice variants and noncanonical genes often prove difficult to annotate.

However, the database size and attendant search space when searching raw genomes, particularly metazoan ones, is usually dramatically inflated. For example, a translation of the human genome in all six reading frames would result in a huge search space: typically thousands of LC-MS/MS spectra would need to be searched against at least $6 \times 10^9$ amino acids. Moreover, eukaryotic genomes are also heavily populated by non-protein-coding regions, introns and hitherto unannotated splice variants. Although it is possible to search directly against translated genomic sequence, this is clearly a challenging task.
that requires careful quality control to avoid false positives and integration of peptide spectrum matches (PSMs) over a genomic locus.

An alternative approach is to search against expressed sequence tags (ESTs) generated from traditional Sanger sequencing, or from next-generation sequencing since they are by definition transcribed (and likely translated) regions of the genome. For example, this approach has been successfully applied to UniGene clusters by generating a six-frame translation and compressing the resulting protein sequence database to reduce redundancy. This allowed the discovery of nonsynonymous mutations, splice-variants, micro exons and alternative translation reading frames, some of which could not be identified from a direct genome search. An alternative method for reducing the search space involves translating the ESTs into proteins using probabilistic approaches such as ESTScan2, DECODER and FrameDP. For example, Robinson and colleagues identified secretory proteins involved in helminth pathogenesis by searching spectra against protein sequences predicted by ESTScan2, which uses a hidden Markov model to distinguish CDS from untranslated regions (UTRs) and correct potential sequencing errors. Such studies have shown that searching against a concise database enriched in protein-coding sequences can increase sensitivity.

Controlling for false positive identifications is essential in high-throughput proteomics studies and particularly so for proteogenomics where intergenic sequence, introns and UTR together can constitute most of the database. Several database search tools such as Mascot, Inspect, X-Tandem or Sequest are commonly used to search spectra against a six-frame translated nucleotide sequence, but the output scores for candidate PSMs cannot be compared directly across experiments. Indeed, it is widely accepted that applying E-value thresholds is anticonservative and different search engines estimate quite different significance levels. Instead, the global error rate is typically estimated by ordering the PSMs according to their E-value or search engine-specific score and calculating the false discovery rate (FDR). The FDR is estimated from the percentage of incorrect PSMs at a given threshold, which is usually calculated using a target-decoy approach. The exact details of this vary and much discussion in the field exists as to the best approach, although the basic principle is well established. This involves searching the spectra against both a target database of ‘real’ sequences and a decoy database containing ‘fake’ sequences produced by reversing or randomizing the target sequences; hits to the latter are considered to be false and are used to estimate the level of incorrect target PSMs at a given score threshold. The FDR applies globally to a collection of PSMs, but individual PSMs can also be associated with a q-value from the FDR level at which they are first reported. The q-value is still a measure of the global error rate within a set of PSMs and like the FDR is dependent on the properties of the database used.

In contrast to the global error rate, a ‘local’ FDR termed the Posterior Error Probability (PEP) is also frequently estimated as the probability of an individual PSM being incorrect. For example, the software tool Quality implements a non-parametric approach for calculating the PEP by estimating the proportion of the target score distribution that is incorrect given a set of p-values or a decoy score distribution. Similarly, PeptideProphet can be used to calculate PEPs and can incorporate decoy search results to enable semiparametric modeling and therefore greater flexibility.

Both the FDR and PEP approaches exploit target/decoy database search results to assign statistical significance to PSMs, presuming that the target database accurately represents genuine protein sequences and the decoy database is of equal (or known) size and similar redundancy. In this study, we examine whether these criteria are indeed met in the context of proteogenomics experiments. Specifically, we highlight the problems associated with the standard target–decoy approach for assessing the statistical significance of PSMs assigned to predicted protein sequences derived from a large collection of Chicken ESTs. We consider searches against six-frame translations and single-frame predicted protein translations, comparing different approaches to estimate the statistical significance of the PSMs. Search results are highly dependent on database choice and suggest potential pitfalls when searching six-frame databases linked to database size and target–decoy error modeling. We believe this leads to overconservative significance estimates for six-frame translation databases and a reduction in sensitivity and, hence, fewer confident PSMs and peptides. We investigate the source of the errors, biases in six-frame translation databases, and suggest a variety of approaches to address these problems. This can be achieved by modifying the six-frame database or by conceptually translating the EST sequences. The results have significance for any group carrying out proteogenomic searches against genomic or transcriptome-based sequences.

**MATERIALS AND METHODS**

**EST Data Set**

A total of 339,314 ESTs were sequenced from 64 cDNA libraries derived from 21 chicken tissues, and then clustered and assembled using BLASTN and PHRAP to generate 85,486 EST contigs. These are available via http://www.chick.ummis.uct.ac.uk. Two different sets of protein sequences were predicted from each EST contig using the EORF and ESTScan2 algorithms.

**Preparation of Chicken Samples and Mass Spectrometry**

We used a comprehensive data set of peptide spectra generated for an unrelated DT40 project, kindly donated by colleagues at the University of Cambridge (Kathryn Lilley, personal communication). The MS/MS data were derived from a proteomic analysis of the DT40 chicken cell line which used the LOPIT protocol, and was originally published in 2009. Full details are available in the original paper, but briefly, DT40 cells were fractionated by density gradient centrifugation and 7 fractions were chosen for analysis. The selected fractions were labeled with four-plex iTRAQ reagents and digested with trypsin. Labeled peptides were pooled together and separated using two-dimensional liquid chromatography. LC-MS/MS was performed using an ultimate-nano-LC system (Dionex) coupled to a QSTAR XL mass spectrometer (Applied Biosystems). The QSTAR XL was operated in information-dependent acquisition mode in which 1 s MS scans were performed (400–1600 m/z) followed by 3 s product ion scans (100–1580 m/z) on the two most intense doubly and triply charged peptides. The LOPIT protocol and iTRAQ labeling were incidental to our study which focuses solely on the relative merits of database composition and attendant statistical treatments to peptide identification.
Figure 1. Schematic of EST translation for target-decoy database generation. Translation of transcriptome data such as ESTs in all six reading frames increases the proportion of ‘junk’ sequence. In this simplified model, only one of the six reading frames is correct (sequence A in frame 2). Sequences denoted by “B” are in the correct direction and therefore in some circumstances could constitute part of the correct ORF as a result of pre-mRNA splicing or frameshift errors. Sequences denoted by “C” are in the wrong direction and are therefore incorrect. Decoy sequences are created by reversing the six corresponding target six sequences, so that decoy1 is the reverse of B1, decoy 2 the reverse of A2, and so on.

Databases

Ensembl protein sequences were downloaded in fasta format from the Ensembl FTP server (ftp://ftp.ensembl.org). Version 56 of the ‘pep.all’ set containing translations from known and novel genes was used. The UniRef90 database (release 15.11) was downloaded from (http://www.uniprot.org/downloads). A custom Perl script was used to generate the six-frame translation database from 85,486 EST contig sequences, generating 3 forward and 3 reverse-frame sequences using the standard genetic code. Protein sequences were also predicted from the EST contigs using the ESTScan2 and EORF programs (manuscript in preparation). Briefly, EORF calculates a score for each reading frame based on its codon usage bias and sequence homology derived from BLAST searches against UniProt, and was used to predict 67,125 ORFs from the EST contigs (a small fraction are rejected as unlikely to be coding). ESTScan2 uses a hidden Markov model to predict the correct reading frame and UTRs. ESTScan2 predicted a total of 62,161 protein sequences from the ESTs (again rejecting noncoding ESTs).

Mass Spectrometry Database Searching

A total of 403,820 centroided spectra were searched against protein sequences derived from the ESTs using Mascot version 2.0, with a precursor MS error tolerance of 0.2 Da and MSMS error tolerance of 0.8 Da. Fixed modifications were Cysteine β-methyliothiolation and iTRAQ labeling of Lysine residues and N-termi. Variable modifications were iTRAQ labeled tyrosine and Methionine oxidation. Up to 1 missed cleavage was permitted.

Five different databases were searched: Ensembl version 56, EORF predictions, ESTScan2 predictions, and the six-frame translations. In addition, various combinations of EORF, ESTScan2 and six-frame sequences were searched to find the database which allows for the most PSMs. Typically, the majority of PSMs for a database search are incorrect; hence, it is important to accurately predict which PSMs represent real peptides present in the samples. For statistical evaluation of the data, decoy databases were constructed by reversing each protein sequence in the original ‘target’ database. Two types of target–decoy searches were performed: separate and composite. In the composite search, the decoy database was concatenated onto the target database and then used as a single database for the Mascot searches, whereas the separate search involved independent searches against target and decoy databases.

Statistical Methods for Validating PSMs

From composite (concatenated) searches, FDRs were calculated using the methods published by Ellias and Gygi and Käll et al. to estimate the proportion of false positive PSMs that have accumulated at a given Mascot score giving FDR$_{lcl}$ and FDR$_{ll}$, respectively. Because the FDR does not increase monotonically with Mascot score, q-values were calculated as the minimum FDR at which a PSM is accepted. The FDRScore method was used to combine PSMs from the EORF and ESTScan2 searches. For each of the two database searches, a custom perl script was used to assign q-values against Mascot scores, to identify step points (where the q-value increases). From this, a linear regression was calculated between each step point. The FDRScore was calculated for each Mascot score between the step points, according to the gradient of the line. PSMs common to both searches were then merged by calculating the geometric mean of their FDRScores. The PSMs were then resorted by their Average FDRScores to calculate a new set of q-values, from which a second FDRScore was calculated, termed the combined FDRScore.

In parallel to the composite database estimates, the Mascot scores for separate database searches were used to calculate the local FDR, or PEP, using the software tool Qvality. Default parameters were used to generate a set of PEPs and PEP-derived q-values linked to Mascot scores.

Estimating the Proportion of Correct PSMs

EST contigs and their attendant reading frames were assigned to Ensembl proteins via BLASTX searches against the Ensembl56 database. Assignments were made for the top
scoring hits which passed the following cut-offs: Identity >95%; Coverage >50 residues; E-value <0.001. EST contigs with significant hits were extracted along with the top scoring reading frame. We assumed that the highest scoring reading-frame contained the correct ORF. This information was integrated with the six-frame translation database, to identify the sequences in the ‘correct’ reading frame owing to the BLASTX match. PSMs having a match with the correct reading frame were then assumed to be correct.

The probability distributions of the amino acid frequencies were calculated for each reading frame and compared with the probability distributions for the Ensembl and UniProt90 protein sequences. The amino acid frequency of the entire correct reading frame set was calculated separately from the incorrect reading frames. The Mann–Whitney U test was used to measure the degree of divergence between the amino acid distributions of the correct frames with each of the incorrect frames different distributions.

## RESULTS AND DISCUSSION

### Searching against Six-Frame or Redundant Databases Affects Sensitivity

When searching high-throughput mass spectrometry data against a protein database using a target–decoy strategy, it is usually the case that the target database is composed of genuine protein sequences that could be present in the sample, and random (false) matches to target and decoy database are equally likely. However, when searching against a six-frame database, these conditions are not necessarily met. One anomaly is that at most only one out of the six possible reading frames translated from a nucleotide sequence is likely to be coding (presuming that there is only one protein coding ORF at any given locus) and can lead to “true” target PSMs. This is illustrated in Figure 1 for a hypothetical EST sequence that codes for a protein, along with its six conceptual translations (forward frames 1, 2, 3 and reverse frames 4, 5, 6) and the attendant six decoy frames. Only the target frame A2 contains a true coding sequence and is matched by a single PSM in this case. Frame B1 and B3 are in the correct direction (and could contain correct PSMs if there were a frameshift mutation in the nucleotide sequence) while frames C4–C6 are not. Although the target and decoy databases contain the same number of sequences, amino acid composition and tryptic peptides, the five “wrong” target frames are not likely to be protein-like. This expansion of the target database is in principle similar to simply adding more proteins (perhaps from another species) to the database, but in this case, the additional targets are clearly “wrong” and not likely to be protein-like in composition. Moreover, the expansion is particularly big, adding five extra sequences for every original target. This could confound the assumptions normally held for target-decoy FDR calculations and lead to incorrect statistical modeling.

To test this, we searched tandem mass spectra derived from a chicken DT40 cell line against different protein databases generated from assembled EST contigs, applying a variety of FDR and PEP-based confidence measures to generate significant PSMs. For FDR estimation, two widely used approaches using concatenated target–decoy databases were used based on methods published by Käll et al. and Elias and Gygi to generate FDRval and FDRperc estimates with attendant q-values. Finally, the PEP and PEP-derived q-values were also estimated from separate database searches using Qvality. As described in the methods, five different protein sequences databases were generated from the EST contigs and assessed by the number of PSMs accepted at a consistent threshold (q-value/PEP < 0.01) for the different confidence measures (Table 1). Decoy sequences were created by reversing the target sequences (see also Materials and Methods).

<table>
<thead>
<tr>
<th>Table 1. Unique Peptide Identifications at q-Value/PEP &lt; 0.01 for Different Database Searches</th>
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<td>Database</td>
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<tr>
<td>ESTScan2</td>
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<td>EORF</td>
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<td>EORF + ESTScan2 + six-frame</td>
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<td>EORF + ESTScan2 (FDRScore)</td>
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Table 1 shows that considerably fewer PSMs were accepted when searching the six-frame database at a q-value/PEP cutoff of 0.01, in comparison to the other databases searched. This observation is independent of the chosen measure of statistical significance, suggesting it is a problem specific to the six-frame database searches (whether concatenated or separate). This reduced sensitivity can be attributed to several factors. First, the six-frame database is more than seven times greater in size than the other databases, and it is known that large databases can lead to reduced search sensitivity owing to more conservative statistical estimates (or conversely, small database size leads to overestimation of significance and likely false positives). Second, the unusual nature of the six-frame database could be confounding and inflating the FDR. In all likelihood, only one of the six forward frames is “correct” and can be matched by genuine peptide spectra. This could lead to an effective imbalance in the true “target” sequences and false “decoy” sequences, since at least five of the target frames are also likely to be wrong, which in turn compromises FDR estimates.

A further possible explanation for the poor PSM sensitivity observed for the six-frame searches could be derived from the ESTScan2/EORF translations correcting errors such as frameshifts, leading to peptides that are absent from the six-frame database. To test this, using the FDRval metric, we compared unique peptide sequences from ESTScan2 PSMs to those derived from six-frame PSMs, in Figure 2a. Only 389 peptides (~3% of ESTScan2 peptides derived from PSMs with q-values less than the 0.01 threshold) are unique to the ESTScan2 database searches. Most of these are expected to come from the translational corrections applied by ESTScan2. While these peptides do contribute to the higher sensitivity of the ESTScan2 search, their overall contribution is minimal. In contrast, 11 902 peptide sequences derive from PSMs found in both databases and 7917 of these have PSM q-values less than the 0.01 significance threshold. Hence, 3985 peptides are “lost” in the six-frame search despite having PSMs in the Mascot output with the same score but with q-values >0.01. Thus, the majority (91%) of the additional ESTScan2 peptides have arisen as a result of the inflated q-values associated with the six-frame target–decoy error modeling as opposed to correction of frame shift errors. It should also be noted that some 703 peptides are exclusive to the six-frame search as ESTScan2 did
not predict a coding sequence, and thus could not be matched by any spectra. These general trends are matched when considering data at the PSM instead of the peptide level (see Supporting Information Figure S1).

In contrast, fewer peptide identifications are “lost” when comparing ESTScan2 and EORF search results, shown in Figure 2b. In this case, ESTScan2 generates more accepted PSMs and attendant unique peptides, but only 1040 additional peptides are “lost” to EORF owing to different FDR modeling while 10810 are shared. The same trend is observed at the PSM level, where only 2% of accepted ESTScan2 PSMs were missed by EORF due to FDR differences, and 87% of accepted ESTScan2 PSMs were also accepted by EORF (see Supporting Information Figure S1).

We also note that combining the EORF and ESTScan2 databases leads to fewer PSMs at an equivalent q-value/PEP threshold, as shown in Table 1. This could be due to increased size as well as redundancy when combining similar databases (i.e., EORF and ESTScan2) where target protein sequences are more likely to be shared. Redundancy of target sequences is already known to degrade search quality; this can be addressed by merging highly similar sequences into a single entry and appending sequence variants in a compatible fashion for MS database searches.22 One potential downside of this approach, however, is the potential loss of subtle variations that represent different biological entities that could be important for proteogenomics.

To circumnavigate this redundancy issue and integrate PSMs from separate ESTScan2 and EORF searches, we used the FDRScore method52 that was initially designed to merge results from multiple search engines.52,54,55 This approach removes redundancy at the PSM level by combining FDRcal values from matched PSMs across two database searches and models database-specific PSMs independently, to derive an integrated FDR estimate. This resulted in 1522 additional PSMs compared to ESTScan2 alone and 5193 compared to the six-frame database. For our spectral data set, this results in the largest number of significant PSMs from all the approaches considered at the same nominal significance threshold of 0.01. However, this is only applicable to the FDRScore method when used to calculate q-values for concatenated target−decoy databases.

The relationship between q-value/PEP and Mascot score was further investigated in order to explain the poor sensitivity of the six-frame translation searches. The PEP, PEP-derived q-values, and empirical q-values (FDRcal) were plotted against Mascot scores for each database search (Figure 3). Figure 3a confirms that the Mascot ion score for equivalent PSMs is independent of the database searched, a relationship that holds for all the paired searches we ran and supports comparisons across the different database searches. Figure 3 confirms that the PEP is clearly the more conservative approach, consistent with previous studies36,44 and that, independent of the significance measure used, the six-frame translation PSMs have higher q-values/PEPs compared to the other database searches.

Biases in the target−decoy database construction methods for FDR calculations have been noted before, which can lead to effective decoy database sizes larger than the target.44 We contend that standard six-frame translation databases may also suffer from related problems leading to inflated q-values and PEPs, as observed in Figure 3. The extent of this inflation can be gauged by comparing the maximum q-values of the different database searches. For example, Figure 3d shows that the maximum PEP-derived q-value (the y-intercept) for the six-frame translation PSMs is much greater than the equivalent q-values for the EORF and ESTScan2 PSMs: approximately 0.8 compared to 0.53. Similarly, the maximum FDRcal q-value for the six-frame database search is much larger than the equivalent q-values from other searches (Figure 3b), and the same trend is also observed for the PEP (Figure 3c). The differences in these profiles show the effect that the choice of error model and database has on search results for high-throughput proteomics.

Collectively, these results point to the statistical modeling and not the search engine scoring that lead to a sensitivity reduction in accepted PSMs at a fixed statistical threshold for six-frame searches, and highlight the importance of redundancy removal and careful database design for estimating PSM statistics in proteogenomics. However, although we observe a broad range of FDR and PEP estimates for the same PSMs, we still do not know which ones are closest to the truth, although one might presume the six-frame values are overly conservative since this is the least sensitive search. We try to address this in the next section.

Figure 2. Overlap of peptides identified in pairwise database searches. Overlap of unique peptide sequences derived from PSMs in the searches against: (a) the ESTScan2 and six-frame databases, (b) ESTScan2 and EORF databases. In both cases, FDRcal q-value cut-offs of 0.01 for the various searches are indicated by dotted lines, black for ESTScan2 and white for the six-frame or EORF searches. PSMs are sorted by Mascot score from low scores (bottom) to high scores (top). The majority of the unique accepted peptides identified in ESTScan2 but missed by the six-frame database were present on both databases but have q-values that exceed the threshold six-frame q-value threshold.
To further study the most appropriate database and error model for the six-frame search results, we investigated an alternative measure of truth for PSMs passing a given statistical threshold. We reasoned that for all EST contigs with significant BLASTX hits to the Chicken Ensembl56 set of proteins, only one of the six translated frames is “correct” based on the top-scoring BLASTX hit to Ensembl. Corresponding PSMs to this “correct” frame were also classified as “correct”, all others as “incorrect”. We therefore expect that close to 100% of all PSMs are “correct” at the lowest q-values/PEPs and, correspondingly, target PSMs with very high q-values/PEPs should be false. Accordingly, PSMs with q-values and PEPs close to zero should be in the “correct” reading frame almost 100% of the time, whereas manifestly incorrect PSMs should be in the “correct” reading frame approximately 17% (one-sixth) of the time. Figure 4 shows that the proportion of correct reading frame PSMs is in fact close to 92.5% at the lowest q-values/PEPs. This small discrepancy can be explained by the fact that the reading frame with the most significant BLASTX hit does not always correspond with (and contain) all the true protein sequence. This will be mostly due to sequencing and mis-assembly errors from the ESTs generating “frameshift” errors that push the true coding sequence in to multiple frames. Therefore, we likely underestimate the proportion of correct PSMs by approximately 7% at the lowest (most significant) q-value/PEP thresholds due to these instances. This represents the maximum error for the estimated percentage of true positive PSMs. Figure 4a shows the percentage of correct PSMs from the total number of PSMs accepted for each q-value calculated using both the FDRKall and FDRREG q-values, as well
psms were considered to be correct if the reading frame contained the top-scoring match to an Ensembl protein through a BLASTX search. Plots show: (a) the percentage of psms falling in the correct reading frame that would be expected at a given q-value threshold, presuming the error modeling is accurate (see Supporting Information Excel File for derivation). For example, approximately 58.3% of PSMS should be in the correct frame at a q-value of 0.5, which falls slightly further still to 54% when we factor in the scaled 7% correction described above (Figure 4a). However, the observed percentages are significantly higher than would be expected by chance, at close to 66%, 73%, and 75% of PSMS assigned to the correct reading frame for the FDRKall, Qality, and FDRq50 based q-values, respectively. This highlights the overly conservative nature of the q-value estimates for the six-frame database search. Interestingly, the FDRKall q-value yields percentages that are closest to expectation, suggesting that for this case, at least, it is a more accurate FDR calculation.

The proportion of true positive PSMS at different PEPs was also estimated, again assigning “correct” reading frames via BLASTX searches (Figure 4b). In this case, the percentage “correct” was calculated within 0.01 bins because the PEP, unlike the q-value, is a local measure of significance which can become overly conservative when thresholds are applied to a list of PSMS, as discussed previously by Käll et al. The extent of this conservativeness can be seen in Figure 4a, where the percentage of correct PSMS remains above 80% for PEPs above 0.8. The binned PEP data in Figure 4b is noisy compared to the cumulative data in Figure 4a, but nevertheless shows a similar trend to the q-value data. Again, the PEP overestimates the expected proportion of true positive PSMS and both q-value and PEP suffer from poor accuracy when deployed without correction in six-frame database searches.

Nevertheless, the PEP is a highly informative statistic in proteogenomics since a novel gene or splice-variant might only be identified via a single PSM, and we therefore need to know the likelihood of this being a correct match. Hence, it has assisted the high-throughput identification of genes and is an alternative to heuristics such as the somewhat arbitrary two-peptide rule. However, most tools require a decoy database in order to generate a null model, which allows tools such as Qality to estimate PEPs more accurately. Likewise, PeptideProphet utilizes decoy databases to allow PEP calculations that are free from parametric assumptions to provide accurate PEPs for different search engines and data sets.

Six-Frame Databases Confound the Target–Decoy Assumption

The source of the atypical statistical estimates generated from the six-frame searches is not immediately obvious. Although relative differences in target and decoy databases sizes can alter the number of accepted PSMS at a given FDR threshold, thereby breaking the target–decoy assumption, this does not apply here. Our target and decoy database sizes are matched in all searches, including the six-frame searches. Indeed, any difference between the number of unique target and decoy peptides can be factored into the FDR calculation using a normalization step; this is implemented in the Qality software used here for separate database searches. Regardless of the metric used to estimate significance, the fewest accepted PSMS are always observed in the six-frame database.

The unique feature of the six frame database is that only one in six target sequences is likely to contain the “true” target and most of the target database is inflated by sequences in the “wrong frame”. This could imbalance the distributions of target and decoy PSM scores, which we investigate here.

Normally, a probability distribution of PSM scores would have a tail to the right, corresponding to the correct target PSMS. Figure 5 shows the distributions of reported target and decoy Mascot scores for the top ranked PSMS identified in the EORF, ESTScan2 and six-frame databases for separate target and decoy database searches. It should be noted here that Mascot does not report PSMS when there is no match (presumably when there is no matching precursor ion or too few matching fragment ions). Hence, in this work, statistics are presented only on the reported PSMS. This has a significant effect as, for example, 369 477 target PSMS are reported from six-frame database searches compared to only 181 699 from ESTScan2, although 403 820 spectra are searched in both cases (see also Supporting Information Table S1).

The EORF and ESTScan2 searches (Figure 5b,c) show a clear difference between the reported target and decoy score
distributions, with the expected tail on the right of high-scoring PSMs that are likely to be correct (i.e., a low PEP or FDR). However, this clear distinction is absent for the six-frame database searches in Figure 5a, with median scores of 14.98 and 13.42 respectively for reported target and decoy PSMs. The ESTScan2 median scores for reported target and decoy PSMs are considerably different and better separated, at 19.33 and 10.77, respectively, with similar values for EORF. The increase in the median target PSM score from six-frame database to ESTScan2 appears counterintuitive since more peptides are present in the six-frame database, including the vast majority of those in the ESTScan2/EORF databases. However, as noted above, considerably fewer target PSMs are reported by Mascot for the ESTScan2 database compared to the six-frame one and these are generally of higher score (which leads to an increased median score), as is evident from Figure 5.

For the six-frame searches, the increase in the reported median decoy PSM score stems from the increased number of candidate sequences against which each spectrum can match. Indeed, Mascot reports a greater number of candidate decoy PSMs overall for the six-frame searches (363,236 compared to 117,982/127,415 in ESTScan/EORF searches). This increase in the number of candidate decoy PSMs for the six-frame searches is consistent with the observed increase in the median decoy PSM score.
would lead to more decoy PSMs out-competing their target equivalents in a concatenated database search, potentially generating more false negatives. This is indeed the case when the cumulative frequency distributions of the target and decoy PSM Mascot scores are considered for the six-frame searches in comparison to ESTScan or EORF (see Supporting Information Figure S2), where decoy PSMs are assigned higher scores in the six-frame database searches for both concatenated and separate search strategies.

This inflation of the target database with 'wrong frame' sequences has consequences for FDR estimates. We illustrate this in Table 2 where we consider a search against a hypothetical standard database that produces 1000 true positive target PSMs and 10 decoy PSMs. We then consider an inflated database that contains the same target sequences plus extra sequences (i.e., alternatively translated 'wrong' frames) that are unlikely to be matched as true positives. We presume the same 1000 true PSMs will be returned from a search against this database, but false positive matches (as estimated from decoy PSMs) will likely increase (say to 20), particularly at modest search engine scores. This is because the decoy database is increased while the effective target database (i.e., “hit-able” sequences) remains the same. The resulting FDR estimate is then increased, leading to fewer accepted PSMs at the same fixed cutoff.

This hypothesis is consistent with the data shown in Figure 5a where additional higher scoring decoy PSMs are identified owing to the inflated nature of the six-frame database. Conversely, the presence of the additional five ‘wrong frame’ target sequences increases the chances of assigning spectra to a false positive target. Indeed, roughly twice as many spectra are assigned to target PSMs by Mascot with reportable scores in the six-frame searches. Somewhat counterintuitively this lowers the median target PSM score (see Figure 5) as Mascot reports more lower scoring PSMs, presumably from hits to the five “wrong frame” sequences. Ultimately, this means that the PEP estimates are increased because they depend on the relative heights of the target and decoy distributions. Likewise, the six-frame FDR will be increased because the decoy PSMs, with relatively higher scores, would be encountered sooner when the sorted list of scores is traversed from high to low during the FDR calculation.

This is illustrated when comparing the ratio of target:decoy PSMs in a concatenated search at a fixed Mascot score between six-frame and ESTScan2 searches (see Supporting Information Figure S3). Both distributions converge toward the expected 0.50 for PSMs with low Mascot scores, although a small bias toward target PSMs remains for the six-frame database even at scores below 5. However, generally, the six-frame ratio is lower, particularly at ion scores above 15. This leads to increased numbers of decoy PSMs, which in turn leads to an increased estimate in false target PSMs and an increased FDR estimate.

The same effect is observed when considering the absolute numbers of target and decoy PSMs in the six-frame database compared to a subset of it composed of just the three forward frames (see Supporting Information Figure S4). Although this database is half the size, it contains most of the true target PSMs (since most correct translations are in one of the forward frames). The numbers of reported target PSMs are almost identical between the two searches at all Mascot scores (with an average ratio of 1.05:1), while there are 1.4X more decoy PSMs on average in the six-frame search results.

The observations detailed above support the hypothesis outlined in Table 2, that the inflated nature of the six-frame translation database is the underlying reason for the overestimation of the FDR for the six-frame searches. However, could this simply be a consequence of increased database size? Since database size will affect FDR and PEP estimates, it has been suggested that larger database sizes can reduce the variance in estimates of the number of decoys (false) PSMs, leading to better (more precise) estimates of significance. However, we note that in general this leads to more conservative estimates (see Supporting Information Figure S5), consistent with previous observations.53 To confirm this, we recalculated the q-value for PSMs drawn from a subset of the six-frame database of identical size to ESTScan2, repeating this 1000 times to generate an average FDR profile, shown in Figure 6. This demonstrates how the atypical database composition must also be influencing the FDR calculations, since the six-frame subset q-value profile is more conservative than the ESTScan2 search despite the fact the database sizes are identical. The increased FDR for the six-frame searches is therefore due both to database size and atypical composition. This highlights the need to exercise caution when generating/modifying proteogenomic databases in order to provide more accurate statistical estimates.

![Figure 6. Effect of database size on FDR of the six-frame PSMs. Subsets of sizes equal to the ESTScan2 database were randomly sampled (1000 times) from six-frame database. The mean q-values were calculated from the samples to give an FDR profile with FDRs greater than the ESTScan2 PSMs, but lower than the six-frame PSMs.](image-url)

Table 2. Hypothetical Target and Decoy PSMs Accepted at a Fixed Score in a Standard and Inflated Database

<table>
<thead>
<tr>
<th></th>
<th>target PSMs</th>
<th>decoy PSMs</th>
<th>FDR at fixed score threshold</th>
</tr>
</thead>
<tbody>
<tr>
<td>Standard database</td>
<td>1000 TPs</td>
<td>10 FPs</td>
<td>= 10/1000 = 0.01</td>
</tr>
<tr>
<td>Inflated database (e.g., six-frame)</td>
<td>1000 TPs</td>
<td>20 FPs</td>
<td>= 20/1000 = 0.02</td>
</tr>
</tbody>
</table>

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Equalizing the Target and Decoy Databases Improves the Sensitivity for the Six-Frame Searches

To estimate accurate FDRs/PEPs, the ratio of target to decoy sequences needs to be as close to 1:1 as possible, or at least properly understood and quantified. We reasoned that reducing the six-frame database to a single target and decoy sequence for each EST contig would provide a more realistic and accurate FDR estimate. A naive but simple way to achieve this is by randomly selecting only one from each of the six target reading frames for each EST contig to generate a random-frame database. A second approach is to select the single frame containing the most PSMs to generate the top-hit-frame database. In both cases, a single target sequence is selected from the EST contigs along with its reversed (decoy) sequence. The associated PEPs were then recalculated for these modified databases and compared to the equivalent values from the standard six-frame searches, shown in Figure 7 and Table 3. As expected, this has a marked effect on PEP estimates for equivalent PSMs between the paired searches. Both methods uniformly lower the PEP compared to equivalent six-frame PSMs leading to additional unique peptide matches. Indeed, although the random-frame database has 83% of the original sequences removed, the number of peptides matched is 657 in excess of expectation at a PEP cutoff of 0.01, presuming that only one-sixth of the original PSMs should remain. This is an interesting finding as it shows that choosing a single frame, effectively at random for many EST contigs, improves relative sensitivity even though some “correct” frames will have been removed by chance. This would not, of course, be a viable strategy in practice since fewer PSMs overall are obtained and many genuine PSMs will be lost, but it serves to illustrate the point that less conservative statistical estimates can be achieved by filtering the database.

We tested other approaches to improve sensitivity, exploiting a priori knowledge on the EST contigs. Given that the direction of sequencing is usually known in transcript libraries, we observe that simply retaining just the forward three frames has an advantage (three-frames-forward), producing over a thousand more unique peptides with PEPs of 0.02 or less (Table 3). Similarly, considering homology to predict the mostly likely coding frame improves performance. For the EST contigs with BLASTX hits to Ensembl56, the single frame with the most significant BLASTX hit to Ensembl proteins. Top-hit-frame refers to the subset of the six-frame database where the frame with the most PSMs is selected for each contig.

Table 3. Unique Peptide Identities at Different PEP Cutoffs Derived from Searches over Different Databases

<table>
<thead>
<tr>
<th>Database Type</th>
<th>PEP cutoff</th>
<th>Number of Target Sequences</th>
</tr>
</thead>
<tbody>
<tr>
<td>Six-frame standard</td>
<td>0.01</td>
<td>4951</td>
</tr>
<tr>
<td>Random-frame</td>
<td>1483</td>
<td>1739</td>
</tr>
<tr>
<td>Three-frames-forward</td>
<td>5654</td>
<td>7050</td>
</tr>
<tr>
<td>Top-hit-frame</td>
<td>6624</td>
<td>8330</td>
</tr>
<tr>
<td>Six-frame-predicted</td>
<td>5601</td>
<td>7251</td>
</tr>
<tr>
<td>ESTScan2</td>
<td>7225</td>
<td>8907</td>
</tr>
<tr>
<td>EORF</td>
<td>7146</td>
<td>8782</td>
</tr>
</tbody>
</table>
This highlights a downside to the BLASTX-filtering approach in this instance; only about 22% of the EST contigs have significant BLAST hits whereas EORF and ESTScan2 predict protein sequences for the majority of them. The additional peptide matches from EORF and ESTScan2 searches may come from novel genes or isoforms that are not yet annotated in the Ensembl56 known gene set.

Collectively, these results demonstrate a variety of approaches to apply when considering proteogenomic searches against nucleotide databases such as EST/cDNA or genomic six-frame translations. Minimally, selecting one single candidate reading frame from the six possible leads to superior FDR and PEP estimates, and using empirical evidence to select one of the six frames via BLASTX searches against existing protein databases is better still. Nucleotide sequences that have no significant BLASTX hits still need to be dealt with, but simple strategies here can also be applied. For example, even weak BLASTX hits can be used to suggest the most likely frame that contains some level of coding features. Similarly, statistics such as codon usage can suggest “protein-like” features that point to the mostly likely frame. We note here that the “correct frame” sequences share similar amino acid composition statistics with Ensembl proteins (see Supporting Information Figure S6, Table S1). Other good practice is evident in the literature, such as the prot4EST pipeline which includes a rule-based approach that identifies the longest ORF from the six-frame trans-

lations.58 Similarly, Adamidi et al.12 selected the three longest ORFs from RNA-seq transcripts for database searching. Another approach, which does not require sequence homology, utilizes information about the experimental protocol used for cloning and sequencing the mRNAs. One small-scale proteomics study identified 51 novel seminal fluid proteins by searching 2D-LC-MS data against only the forward frames of translated UniGenes.59 When applied to shotgun proteomics data this improves sensitivity relative to standard six-frame searches, a trend we also note here. Our EST data is generated from directional cloning and subsequent 5′-end sequencing of chicken cDNAs, leading to a bias where 97% of the “correct” frames from BLASTX searching are in frames one, two, or three (see Supporting Information Figure S7). Searching against this three-frame database leads to over 1000 additional peptide identifications filtered at a PEP < 0.02 (Figure 7, Table 3). This improved performance is likely the result of the reduction in erroneous targets, which is reduced to two out of three in the 3-frame search from five out of six. Although the performance is inferior to the BLASTX-filtered single-frame and ESTScan/ EORF searches, the three-frame approach can be used for all EST contigs and not just those with BLASTX matches. Moreover, this approach can be easily implemented for newly sequenced genomes when there are no homologues annotated within the clade, something that would create problems for EORF/ESTScan2 as these tools require codon usage information.

Ultimately, however, our most sensitive searches were derived from the EST translation tools, ESTScan2 and EORF, which generate a single translated sequence across the most likely reading frames. The most effective search strategy overall in terms of sensitivity involved the two database searches using the FDRScore algorithm for combining multiple search results.52

### CONCLUSION

Our study has highlighted some of the pitfalls when searching against nucleotide databases via six-frame translations and how inflating the target database with incorrect sequences perturbs FDR and PEP estimates. Although some methods exist for reducing the peptide search space in proteogenomic studies,9,10 most do not formally consider the inherent biases that can arise through the improper use of the target—decoy approach. We show here that naive six-frame searching leads to overconservative statistics and potential loss of high-quality peptide evidence for genomic annotation. This parallels any shotgun experiment that searches against very large protein databases (such as all of NCBI or UniProt), containing many “unhittable” sequences from other species, which can also lead to overestimated statistics and reduced sensitivity.

Although formally our results have been generated only from assembled ESTs, we see no reason to suppose that the general principles we observe should not apply to RNA-seq data or similar nucleotide sequences such as from gene prediction software or even raw genomic sequence. We recommend proteogenomic practitioners consider the following redundancy removal guidelines when searching against nucleotide databases.

- **Selection of most likely frame based on PSMs.** Retention of the single frame (and its decoy) with most PSMs prior to calculation of FDR/PEP provides less conservative statistics, although multipass approaches need to be performed appropriately.43,60,61
- **Selection of most likely frame based on homology.** BLASTX searches against a protein database, to identify the single frame mostly likely to be coding removes redundancy and improves sensitivity. If close homologues are not available, even weak matches to known proteins should enrich for coding sequence.
- **Selection of frame based on coding potential.** When no homologues are available, similarity in amino acid composition or codon usage can help select the most likely coding frame.
- **Translation software.** Gene prediction (e.g., Augustus) or EST translation software (such as ESTScan or EORF) can overcome introns and frameshifts, as well as remove redundancy. EST translators can be combined with BLASTX data to improve accuracy.58

Ideally, we suggest that gene prediction or EST translation software are likely to be most effective, since six possible frames are reduced to one while frameshift mutations and introns can be overcome. Moreover, presence of a transcript is bone fide evidence that a genomic region is at least transcribed and in most cDNA libraries the sequence is normally largely free of introns and the clonal direction is also known, so only three forward frames need be considered.

One should not lose sight of the raison d’être of proteogenomics, to identify novel genes and novel gene structure that may not exist in extant genome annotations or transcriptome data sets. This necessitates searching against translated genome sequence, and in such cases, use of a gene prediction tool such as Augustus,2,63 perhaps used with reduced stringency to capture some atypical genes/gene structure, is preferable to raw searches against six-frames. Alternatively, if searching against raw genomic sequence we suggest selecting one candidate translation (out of six) with the closest homology with Ensembl or UniRef proteins (i.e., similar
to the six-frame-predicted database), in addition to applying compositional filters to limit regions unlikely to be coding. This should improve the overall sensitivity of the searches while improving the statistical modeling of incorrect PSMs.

It is also worth noting an additional downside to selecting a single reading frame, when several true coding regions overlap at a single genomic locus but only a single one is selected in silico (the INK4/ARF locus is one such example). However, this might be overcome by screening the six-frame database for contigs that contain multiple high-confidence PSMs in different reading frames. Equally, transcriptomic data might very well capture the two independent but overlapping coding regions in different clusters.

Finally, it should be stressed that care needs to be taken when applying target–decoy strategies to shotgun proteomics data and proteogenomics data in particular. As several authors have pointed out, if the underlying database or search engine is not compliant with the assumptions of the target–decoy approach, significance values can be underestimated.41,43,60,61 This is particularly acute for multipass search strategies that change the target–decoy structure between search phases. Here, we ensure the numbers of candidate target and decoy sequences in the database are the same, though we do favor target sequences with certain properties (i.e., those likely to be coding, have the most PSMs, etc.), which could introduce some biases. Nevertheless, the evidence that standard six-frame searches considerably overestimate FDRs/PEPs seems overwhelming, which highlights the need for databases to be checked for compliance with the target–decoy approach (in addition to the search engines). It is clear that more work is needed to produce better error models and databases to develop more reliable statistical confidence estimates, which would greatly benefit proteogenomics by helping to achieve a more comprehensive representation of the proteome.

### ASSOCIATED CONTENT

- Supplementary Figures S1–S7 and Table S1, Venn diagrams of overlap between search results at the PSM level, cumulative frequency plots of mascot ion scores, percentage of target PSMs versus mascot ion score, comparisons of six-frame and three-frame search statistics, scatterplots of accepted peptides versus database size, amino acid composition histograms, a table of frame search statistics, scatterplots of accepted peptides versus mascot ion score, comparisons of six-frame and three-frame approaches, significance values can be underestimated.41,43,60,61
- This is particularly acute for multipass search strategies that change the target–decoy structure between search phases. Here, we ensure the numbers of candidate target and decoy sequences in the database are the same, though we do favor target sequences with certain properties (i.e., those likely to be coding, have the most PSMs, etc.), which could introduce some biases. Nevertheless, the evidence that standard six-frame searches considerably overestimate FDRs/PEPs seems overwhelming, which highlights the need for databases to be checked for compliance with the target–decoy approach (in addition to the search engines). It is clear that more work is needed to produce better error models and databases to develop more reliable statistical confidence estimates, which would greatly benefit proteogenomics by helping to achieve a more comprehensive representation of the proteome.

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