Depression symptoms reduce physical activity in COPD patients: a prospective multicenter study

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Depression symptoms reduce physical activity in COPD patients: a prospective multicenter study

Iván Dueñas-Espín1–5
Heleen Demeyer6
Elena Gimeno-Santos1–3
Michael I Polkey7
Nicholas S Hopkinson7
Roberto A Rabinovich8
Fabienne Dobbels9
Niklas Karlsson10
Thierry Troosters6,11
Judith Garcia-Aymerich1–3

On behalf of the PROactive Consortium

1ISGlobal, Centre for Research in Environmental Epidemiology (CREAL), Universitat Pompeu Fabra (UPF), 2CIBER Epidemiología y Salud Pública (CIBERESP), Barcelona, Spain; 3Secretaria Nacional de Educación Superior, Ciencia, Tecnología e Innovación del Ecuador (SENERCYT), Quito, Ecuador; 4Institut d’investigacions Biomèdiques August Pi i Sunyer (IDIBAPS), Barcelona, Spain; 5Department of Rehabilitation Sciences, KU Leuven, Leuven, Belgium; 6NIHR Respiratory Biomedical Research Unit at the Royal Brompton and Harefield NHS Foundation Trust and Imperial College, London, UK; 7ELEGI Cols Laboratory, Centre for Inflammation Research, The Queen’s Medical Research Institute, University of Edinburgh, Edinburgh, UK; 8Department of Public Health and Primary Care, Academic Centre for Nursing and Midwifery, KU Leuven, Leuven, Belgium; 9Health Economics and Outcomes Research, AstraZeneca R&D, Mölndal, Sweden; 10Department of Respiratory Diseases, University Hospitals Leuven, Leuven, Belgium

Correspondence: Judith García-Aymerich
Centre for Research in Environmental Epidemiology, Doctor Aiguader 88, 08003 Barcelona, Spain
Tel +34 93 214 7350
Fax +34 93 214 7302
Email jgarcia@creal.ca

Background: The role of anxiety and depression in the physical activity (PA) of patients with COPD is controversial. We prospectively assessed the effect of symptoms of anxiety and depression on PA in COPD patients.

Methods: We evaluated anxiety and depression (Hospital Anxiety and Depression Scale [HADS]), PA (Dynaport® accelerometer), and other relevant characteristics in 220 COPD patients from five European countries at baseline and at 6 and 12 months of follow-up. HADS score was categorized as: no symptoms (score 0–7), suggested (8–10), and probable (>11) anxiety or depression. We estimated the association between anxiety and depression at t (baseline and 6 months) and PA at t+1 (6 and 12 months) using regression models with a repeated measures approach.

Results: Patients had a mean (standard deviation) age of 67 (8) years, forced expiratory volume in 1 second 57 (20)% predicted. At baseline, the prevalence of probable anxiety and depression was 10% and 5%, respectively. In multivariable models adjusted by confounders and previous PA, patients performed 81 fewer steps/day (95% confidence interval, −149 to −12, P=0.02) per extra point in HADS-depression score. HADS-anxiety symptoms were not associated with PA.

Conclusion: In COPD patients, symptoms of depression are prospectively associated with a measurable reduction in PA 6 months later.

Keywords: COPD, anxiety, depression, HADS, physical activity, prospective study

Introduction

The determinants of physical activity (PA) in patients with COPD are poorly understood. A recent systematic review identified several clinical, functional, socio-demographic, and lifestyle factors associated with PA in these patients, but the quality of evidence was low to very low, mainly due to the use of a cross-sectional design and lack of control for confounders. This lack of information is in contrast with the importance of PA in COPD, as regular PA has been consistently related to a reduced risk of hospitalizations and death.1

Among the potential determinants of PA that can be modified at the clinical setting, anxiety and depression are relevant because they are highly prevalent in COPD patients and affect COPD prognosis.2–4 Anxiety has been reported in ~40% of COPD patients.5 It can lead to tachypnea, which may worsen lung hyperinflation, leading to increased exertional dyspnea, reduced exercise capacity, and poor quality of life.6 Depression, in turn, has been identified in ~25% of COPD patients7 and is also associated with poorer exercise capacity and worse health status.2

Previous studies have assessed the association between anxiety and depression symptoms with PA.7–15 However, they have a cross-sectional design so the directionality of the association (eg, whether anxiety/depression reduces PA or...
PA reduces anxious/depressive symptoms) could not be established. Moreover, prior analyses have not considered the role of potentially relevant confounders, like exercise capacity, or smoking status. Hence, it remains unclear if anxiety or depression exerts an independent effect on PA levels alongside other factors commonly influencing these levels. Finally, most studies mentioned used indirect measures of PA such as questionnaires, so there is potential for misclassification of the primary outcome of interest (ie, PA).

We therefore aimed to assess the effect of anxiety and depression symptoms on PA in patients with COPD, overcoming the limitations of previous studies by using a prospective design, measuring PA with an accelerometer and controlling for a wide range of potential confounders. We hypothesized that symptoms of anxiety and depression are inversely related with future PA in COPD patients. This study is part of the “Physical Activity as a Crucial Patient Reported Outcome in COPD (PROactive)” project.

Methods
See complete version in the Supplementary material.

Design
Prospective study, with repeated measures at baseline, 6 months, and 12 months. Clinical trial registration: www.clinicaltrials.gov; number: NCT01388218.

Study population
Between July 12th and November 18th, 2011, we recruited a consecutive sample of 236 patients with COPD defined by spirometry (post-bronchodilator forced expiratory volume in 1 second to forced vital capacity ratio [FEV₁/FVC] <0.70), and followed them for 1 year with visits at baseline and at 6 and 12 months. Patients were recruited from tertiary hospitals, rehabilitation centers, and primary care settings from Athens (Greece), Edinburgh and London (UK), Leuven (Belgium), and Groningen (the Netherlands). Inclusion criteria were: i) men or women at least 40 years of age, ii) with medical diagnosis of COPD (FEV₁/FVC less than 70%) confirmed by spirometry, iii) smoking history equivalent to at least 10 pack years, and iv) ability to read and write and to use electronic devices and PA monitors. Exclusion criteria were: i) orthopedic, neurological or other complaints that significantly impair normal biomechanical movement patterns, ii) respiratory diseases other than COPD (eg, asthma), and iii) cognitive impairment; further details on inclusion and exclusion criteria have been reported elsewhere.

For the current study we additionally excluded patients who at baseline had a clinical diagnosis and/or active treatment for anxiety or depression (n=16). Thirty-seven patients were lost to follow-up; there were no baseline differences between patients followed and lost to follow-up except for a lower frequency of cardiovascular diseases in those lost to follow-up (32.4% vs 57.9%, P-value=0.05) (Table S1). No patients were diagnosed with anxiety or depression during follow-up.

This study was conducted in accordance with the amended Declaration of Helsinki, advised and approved by the PROactive ethics and patient advisory boards, approved by local institutional review boards (Commissie Medische Ethiek van de Universitaire Ziekenhuizen KU Leuven; Medische Ethische Toetsingcommissie Universitair Medisch Centrum Groningen; Lothian Ethics Committee; South East Scotland Research Ethics Committee; Scientific Council of the General Hospital for Chest Diseases, Sotiria), and written informed consent was obtained from all patients.

Measurements
At baseline, we collected age, sex, ethnicity, marital status, living status, socio-economic status, and working status. Anxiety and depression symptoms were evaluated at baseline and at 6 months, using the linguistically and culturally validated versions of the Hospital Anxiety and Depression Scale (HADS) that provides a HADS-depression (HADS-D) and a HADS-anxiety (HADS-A) score. Both total scores range from 0 to 21. According to literature in the field values from 0 to 7 are considered normal, a score of 8 to 10 is suggestive of the presence of the corresponding mood disorder, and a score of 11 or higher indicates probable presence of the corresponding mood disorder.

PA was measured at baseline, and at 6 and 12 months (mean [standard deviation, SD]) lapse 177 (12) days between baseline and 6 months visits, and 179 (14) days between 6 and 12 months visit, using a validated accelerometer, the Dynaport MoveMonitor (McRoberts BV, The Hague, the Netherlands). Patients were instructed to wear it at least 10 hours per day during 7 days. We defined an a priori valid activity monitor measurement if patients recorded at least 3 days per week (not necessarily consecutive) with at least 10 hours per day, without distinction between weekdays and weekends, according to previous literature. We obtained the following variables: wearing time, steps per day, and time in locomotion.

We collected additional information at baseline and at 6 months as detailed in the Supplementary material, briefly, smoking history, COPD exacerbations history, dyspnea, quality of life (COPD assessment test [CAT]), doctor’s diagnosis of comorbidities, treatments, body mass index (BMI), fat free
mass index, lung function (FEV₁, FVC, total lung capacity [TLC], residual volume [RV], RV/TLC), and exercise capacity (6 minutes walking distance test [6MWD]).

**Statistical analysis**

The number of subjects available was greater than required according to sample size calculations (see Supplementary material). We assessed the presence and patterns of missing values and, assuming the missing-at-random hypothesis (ie, missingness conditional on measured patients’ characteristics) (see Supplementary material for further detail), we used multiple imputation (100 imputed datasets) with the method of chained equations. We compared the characteristics of patients in the complete case to the imputed datasets (Table S2). All analyses were performed using the imputed dataset.

In order to maximize data use we used a repeated measures approach (Figure 1), each patient contributing twice to the analysis. Let \( t \) denote each time period (baseline, 6 months, and 12 months). For each patient, we estimated the association between anxiety or depression at \( t \) and PA at \( t+1 \), ie, anxiety or depression at baseline and PA at 6 months, and anxiety or depression at 6 months and PA at 12 months, respectively. In bivariate and multivariable analyses we accounted for the correlation between observations of the same subject.

First, we tested the bivariate associations between anxiety and depression (both categorized according to clinically relevant cut-offs and as continuous variables) and PA (steps/day and time in locomotion), using analysis of variance and linear regression, given that PA outcomes were normally distributed. Prior to building multivariable models we used generalized additive model analyses to test whether the shape of the association between anxiety or depression and PA was linear. Then, we used multivariable linear regression to assess the independent effect of anxiety and depression on PA. We tested goodness-of-fit of the models according to sex, health-related quality of life, COPD spirometric severity, smoking status, comorbidities, exercise capacity, socio-economic status, and BMI (Supplementary material). To test the potential role of COPD exacerbations during follow-up as mediators of the effect of anxiety or depression on PA, we repeated the analyses: i) including “COPD exacerbations during follow-up” as a covariate in multivariable models, and ii) excluding patients who had at least one COPD exacerbation during the follow-up period.

We performed several secondary analyses (Supplementary material) to assess the sensitivity of our estimates to our assumptions regarding biases, as well as to test for model misspecification. Briefly, we repeated analyses: i) using PA at 12 months only (ignoring measures at 6 months), ii) additionally adjusting for season of PA assessment, and iii) using complete case dataset. The analyses were performed using Stata 12.1 (Stata Statistical Software: Release 12; StataCorp LP, College Station, TX, USA) and R 3.0.1 (R Foundation for Statistical Computing, Vienna, Austria).

**Results**

A total of 220 patients were included, distributed similarly across study sites (48 [21.8%] from Athens, 41 [18.6%] from Edinburgh, 48 [21.8%] from Leuven, 48 [21.8%] from London, and 35 [15.9%] from Groningen). A majority of patients were men (68%), with a mean (SD) age of 67 (8) years, FEV₁ 57 (20)% predicted, RV/TLC 51 (11)%; BMI 27.1 (5.5) kg/m², and a moderate quality of life (Table 1). The baseline prevalence of suggested and probable anxiety was 19% and 10% respectively, and the corresponding prevalence for depression was 16% and 5%. All followed patients had valid (ie, at least 3 days with at least 10 hours wearing time) data from the accelerometer. The mean (SD) Dynaport wearing time was 880 (120) min/day, the mean number of steps was 4,812 (3,147) and mean time in locomotion was 60 (36) min/day. During follow-up, 22 (10%) patients were hospitalized at least once for a COPD exacerbation.

![Figure 1](https://www.dovepress.com/1289)

**Figure 1** Approach to assess the effect of anxiety and depression (exposures) at \( t \) and physical activity (outcome) at \( t+1 \) adjusting for covariates and physical activity at \( t \).
Table 1 Baseline sociodemographic, clinical, and functional characteristics of 220 COPD patients from five European sites

<table>
<thead>
<tr>
<th>Baseline characteristics</th>
<th>All patients n=220</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (years), m (SD)</td>
<td>67 (8)</td>
</tr>
<tr>
<td>Sex: male, n (%)</td>
<td>149 (68%)</td>
</tr>
<tr>
<td>Living alone, n (%)</td>
<td>108 (49%)</td>
</tr>
</tbody>
</table>
| Socio-economic status:  
  low (IV and V), n (%)   | 30 (14%)           |
| Active worker, n (%)     | 26 (12%)           |
| Smoking status: current, n (%) | 38 (17%) |
| Pack years, median (P25 to P75) | 47 (31 to 72) |
| Any COPD exacerbation in the previous year, n (%) | 88 (40%) |
| mMRC score, median (P25 to P75) | 2 (1 to 2) |
| Number of chronic diseases (other than COPD), median (P25 to P75) | 2 (1 to 3) |
| Cardiovascular disease, n (%) | 118 (54%) |
| Diabetes, n (%)          | 19 (9%)            |
| Musculoskeletal diseases, n (%) | 48 (22%) |
| Osteoporosis, n (%)      | 23 (10%)           |
| Body mass index (kg/m²), m (SD) | 27 (5) |
| Fat free mass index (kg/m²), m (SD) | 18 (3) |
| FEV₁ (% predicted), m (SD) | 57 (20) |
| COPD severity stages     |                    |
| I – Mild (FEV₁ >80%)     | 30 (14)            |
| II – Moderate (FEV₁ <80% and ≥50%) | 101 (46) |
| III – Severe (FEV₁ <50% and ≥30%) | 69 (31) |
| IV – Very severe (FEV₁ <30%) | 20 (9) |
| RV/TLC (%), m (SD)       | 51 (11)            |
| Six minutes walking distance test, m (SD) | 423 (130) |
| CAT score (0 to 40), median (P25 to P75) | 14 (8 to 21) |
| Probable anxiety (HADS-A ≥11), n (%) | 23 (10%) |
| Probable depression (HADS-D ≥11), n (%) | 12 (5%) |

Notes: Socio-economic status was measured through the UK National Statistics Socio-economic classification. Cardiovascular disease was the presence of doctor-diagnosed diseases in medical records (ICD9 390–459). Musculoskeletal diseases included connective tissue diseases (ICD9 710–739). Osteoporosis included osteoporosis (ICD9 733.0) and osteopenia (ICD9 733.90).

Abbreviations: m, mean; SD, standard deviation; P25, 25th percentile; P75, 75th percentile; mMRC, Modified Medical Research Council Dyspnea Scale; FEV₁, forced expiratory volume in 1 second; RV/TLC, residual volume to total lung capacity ratio; CAT, COPD assessment test; HADS-A, Hospital Anxiety and Depression Scale-Anxiety; HADS-D, Hospital Anxiety and Depression Scale-Depression; ICD, International Classification of Diseases.

In bivariate analysis, depression symptoms at t were associated with a lower amount of PA at t+1 (6 months later) in a linear dose-response shape: 4,752 (2,334), 3,837 (1,987), and 3,446 (2,144) steps/day in patients with no symptoms, suggested and probable depression, respectively, P-trend =0.02 (Figure 2 and Table S3). Corresponding time in locomotion was 59.2 (37.3), 49.1 (31.5), and 39.8 (44.1) min/day, P-trend =0.01. The differences of PA across HADS-A categories were not statistically significant.

In the multivariable saturated model (adjusted for age, 6MWD, comorbidities, FEV₁, and steps at t), patients walked 70 steps/day less (P=0.05) for each extra point on the HADS-D score (Table 2). This association persisted in the parsimonious model (β=-81 steps/day, P=0.02). HADS-D score was also significantly associated with less time in locomotion (Table S4). HADS-A score was not statistically associated with any PA outcome in multivariable models.

The effect of HADS-D on steps did not change after stratifying by potential effect modifiers, with two exceptions: i) a stronger association among patients with moderate CAT scores (β=-132, P=0.03) compared to patients with lower or higher CAT scores (β=-19, P=0.79 and β=-26, P=0.73, respectively), and ii) a stronger association among people with high socio-economic status compared to those with low socio-economic status (β=-88, P=0.02, vs −2, P=0.98, respectively) (Table S5).

The inclusion of COPD exacerbations during follow-up in the model did not modify the estimate of the association between depression and steps either (β=-70 steps/day per each HADS-D point, P=0.05) (Table S6). Similarly, the exclusion of patients with any COPD exacerbation during follow up (n=101, 46%) did not modify such estimate (β=-81, P=0.15) (Table S6). Sensitivity analyses yielded very similar results (Tables S7–S9). Linear regression goodness-of-fit tests did not reveal any abnormality.

Discussion

Our study shows that depression symptoms, as measured by the HADS, are prospectively associated with fewer steps per day and less time in locomotion in COPD patients after 6 months. However, no effect of anxiety on PA variables was found.

Previous studies reported a cross-sectional association between depression and PA levels in COPD patients. Our findings move beyond these preliminary findings by testing the directionality of this association. Indeed, presence of depressive symptomatology very quickly seems to negatively impact PA (ie, over a time period of 6 months already), even after adjusting for baseline PA levels. Furthermore, in contrast to previous studies, our study for the first time controlled for a comprehensive set of potential confounders such as lung function, exercise capacity or comorbidities. Moreover, while previous studies focused on the role of depressive disorder, we demonstrated that presence of depressive symptoms impacts PA, a finding that is of major clinical importance. As each extra point in the HADS-D score was associated with a reduction of 81 steps, our results support that COPD management includes the assessment and treatment of mood disorders, even at subclinical stages. We would suggest that further research includes randomized controlled trials of drug and non-drug interventions to

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reduce depression symptoms, since they may be valuable also to improve PA.

Several mechanisms have been proposed that support the plausibility of depression symptoms affecting PA in both COPD and healthy individuals. First, a recent systematic review found a clear association between depression and hospital admissions caused by COPD exacerbations and an effect of exacerbations on PA levels has also been reported. However, we could not confirm that the effect of depression on PA is mediated by COPD exacerbations. Second, hormonal changes that appear during depression, including over-activity of the hypothalmo-pituitary-adrenocortical axis or dysregulation of the autonomous nervous system, have been associated with reduced PA in experimental studies and deserve further scrutiny in COPD patients. Finally, symptoms of depression may exacerbate the perception of dyspnea, which in turn could reduce PA levels.

The lack of a measurable confounding role of exercise capacity in the association between depression and PA in our study, together with the large body of evidence showing a weak to moderate correlation between exercise capacity and daily PA suggests that the clinical effects of depression on COPD patients may be underestimated if only controlled assessment of exercise capacity is performed at the clinical

Figure 2 Distribution of steps and time in locomotion at t+1, according to categories of HADS-A and HADS-D at t.

Notes: (A) steps at t+1 across HADS-A categories at t; (B) steps at t+1 across HADS-D categories at t; (C) time in locomotion t+1 months across HADS-A categories at t; and (D) time in locomotion at t+1 across HADS-D categories at t. HADS-A score: 0–7 normal, 8–10 suggested anxiety, and ≥11 probable anxiety. HADS-D score: 0–7 normal, 8–10 suggested depression, and ≥11 probable depression.

Abbreviations: HADS-A, Hospital Anxiety and Depression Scale-anxiety; HADS-D, Hospital Anxiety and Depression Scale-Depression.
or research setting. Further, we found a stronger effect of depression on PA among people with moderate impact on quality of life than in those with low or high impact on quality of life. It could be speculated that patients with a moderate quality of life are those who experience the bigger impact of depressive symptoms on health outcomes, because in milder stages the depression symptoms are lower while in worsened quality of life, patients have reached the acceptance of the disease together with a higher understanding of their limitations. As previous research has reported a strong relationship between depression symptoms and quality of life scores in COPD patients, it is likely that the association between depression and PA

The strengths of our study include the exhaustive control for potential confounders, the use of direct methods to assess PA, and the longitudinal design. Limitations of our study are the following. Information bias in the measure of anxiety and depression is possible because the questionnaire is subject to potential mis-classification. However, HADS has been demonstrated to have good psychometric properties. It could be argued that the proportion of patients with HADS anxiety and HADS depression $\geq 11$ was low (10% and 5%, respectively) in comparison with other studies. This can be due to the fact that we excluded patients with a diagnosis and/or treatment for clinical depression and/or anxiety (to avoid any effect of therapies on PA level). Consequently, it is likely that the association between depression and PA level is higher in the whole COPD population than in our sample. In any case, confirmation of our results in studies using other validated instruments to assess anxiety and depression is warranted. Secondly, it could be argued that our population does not represent the full COPD spectrum and so the results cannot be generalized to all patients. However, it is worth noting that this risk was mitigated by including patients recruited in diverse settings (primary care, hospital, and rehabilitation) and a range of mild to very severe airflow limitation. Anyhow, any difference in the prevalence of depression, anxiety or PA levels with respect to the full COPD population would not have affected the association between them. Third, reverse causation, ie, that PA affects depression symptoms, could have been an alternative interpretation if the study design was cross-sectional. However, our study had a longitudinal prospective design, with anxiety

<table>
<thead>
<tr>
<th>Characteristics at t</th>
<th>Crude$^a$</th>
<th>Adjusted Saturated model$^b$</th>
<th>Parsimonious model$^c$</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>$\beta$ (95% CI)</td>
<td>P-value</td>
<td>$\beta$ (95% CI)</td>
</tr>
<tr>
<td>Constant</td>
<td>N/A</td>
<td>–</td>
<td>4,690 (4,323 to 5,058)</td>
</tr>
<tr>
<td>HADS-D (per each extra point)</td>
<td>$-204 (-313 to -96)$</td>
<td>$&lt;0.01$</td>
<td>$-70 (-141 to 0)$</td>
</tr>
<tr>
<td>Age (per each extra year old)</td>
<td>$-86 (-131 to -42)$</td>
<td>$&lt;0.01$</td>
<td>$-15 (-50 to 19)$</td>
</tr>
<tr>
<td>Six minutes walking distance test (6MWD, per each extra meter)</td>
<td>11 (8 to 13)</td>
<td>$&lt;0.01$</td>
<td>1.7 (0.8 to 4.2)</td>
</tr>
<tr>
<td>$\geq 2$ comorbidities</td>
<td>$-904 (-1,785 to -182)$</td>
<td>0.01</td>
<td>$-332 (-833 to 169)$</td>
</tr>
<tr>
<td>FEV$_1$, % predicted (per each extra percent unit)</td>
<td>42 (21 to 63)</td>
<td>$&lt;0.01$</td>
<td>10 (-3 to 22)</td>
</tr>
<tr>
<td>Steps/day at t (per each extra step)</td>
<td>0.7 (0.6 to 0.8)</td>
<td>$&lt;0.01$</td>
<td>0.6 (0.5 to 0.8)</td>
</tr>
</tbody>
</table>

Notes: *Each row is a single model. The column is a single model. The value of the constant is the number of steps at t+1 of a patient with mean HADS-D score, mean age, mean 6MWD, $<2$ comorbidities, mean FEV$_1$, % predicted and mean steps/day at t. The column is a single model. The value of the constant is the number of steps at t+1 of a patient with mean HADS-D score, $<2$ comorbidities, mean FEV$_1$, % predicted and mean baseline steps/day at t.

Abbreviations: CI, confidence interval; N/A, not applicable; HADS-D, Hospital Anxiety and Depression Scale-Depression; FEV$_1$, forced expiratory volume in 1 second.

Table 2 Crude and adjusted associations between HADS-depression score at t and daily steps at t+1 (linear regression model)
and depression being measured at baseline, PA measured 6 months later, and baseline PA having been included as a covariate in the regression models to account for PA history and potential reverse causation. Finally, although the possibility of residual confounding can never be ruled out, to the authors’ knowledge, all potential confounders were collected and tested.

Conclusion
In conclusion, symptoms of depression are prospectively associated with a measurable reduction in PA 6 months later in COPD patients. Whereupon, and since the current management guidelines for patients with COPD include the treatment of comorbidities, our study suggests that the management of symptoms of depression need to be considered, including in subclinical states, to prevent decline in PA over time.

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Author contributions
JGA takes responsibility for (is the guarantor of) the content of the manuscript, including the data and analysis. IDE and JGA led the study, performed the statistical analysis, and drafted the manuscript. All authors contributed to conception, hypotheses delineation, design of the study, acquisition of the data or analysis and interpretation of results; critically revising the article; approved the final version to be published; and agree to be accountable for all aspects of the work.

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The effects of anxiety and depression on physical activity in COPD