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A Randomized, Single-Blind, Crossover Trial of Recovery Time in High-Flux Hemodialysis and Hemodiafiltration

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Bernard G. Francq, PhD,3 Alex McConnachie, PhD,3 and Robert Mactier, MD, FRCP1

Background: The choice between hemodiafiltration (HDF) or high-flux hemodialysis (HD) to treat end-stage kidney disease remains a matter of debate. The duration of recovery time after treatment has been associated with mortality, affects quality of life, and may therefore be important in informing patient choice. We aimed to establish whether recovery time is influenced by treatment with HDF or HD.

Study Design: Randomized patient-blinded crossover trial.

Settings & Participants: 100 patients with end-stage kidney disease were enrolled from 2 satellite dialysis units in Glasgow, United Kingdom.

Intervention: 8 weeks of HD followed by 8 weeks of online postdilution HDF or vice versa.

Outcomes: Posttreatment recovery time, symptomatic hypotension events, dialysis circuit clotting events, and biochemical parameters.

Measurements: Patient-reported recovery time in minutes, incidence of adverse events during treatments, hematology and biochemistry results, quality-of-life questionnaire.

Results: There was no overall difference in recovery time between treatments (medians for HDF vs HD of 47.5 [IQR, 0-240] vs 30 [IQR, 0-210] minutes, respectively; \(P = 0.9\)). During HDF treatment, there were significant increases in rates of symptomatic hypotension (8.0% in HDF vs 5.3% in HD; relative risk [RR], 1.52; 95% CI, 1.2-1.9; \(P < 0.001\)) and intradialytic tendency to clotting (1.8% in HDF vs 0.7% in HD; RR, 2.7; 95% CI, 1.5-5.0; \(P = 0.002\)). Serum albumin level was significantly lower during HDF (3.2 vs 3.3 g/dL; \(P < 0.001\)).

Health-related quality-of-life scores were equivalent.

Limitations: Single center; mean achieved HDF convection volume, 20.6 L.

Conclusions: Patients blinded to whether they were receiving HD or HDF in a randomized controlled crossover study reported similar posttreatment recovery times and health-related quality-of-life scores.

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INDEX WORDS: Hemodialysis (HD); high-flux HD; hemodiafiltration (HDF); recovery time; intradialytic hypotension; symptomatic hypotension; blood pressure; dialysis circuit clotting; dialysis modality; end-stage kidney disease (ESKD); quality of life; randomized controlled trial (RCT).
scores than the general population, and this is associ-
ated with increased morbidity and mortality.\textsuperscript{16,17} Although some previous studies have shown an
improvement in HRQoL with convective treatments
compared to HD,\textsuperscript{3,4} the largest RCT to have studied this
outcome found no difference between HDF and low-
flux HD.\textsuperscript{6} Length of recovery time after dialysis is an
important patient-reported outcome measure that
adversely affects HRQoL, and evidence from the
Dialysis Outcomes and Practice Patterns Study
(DOPPS) cohort has suggested an association between
longer postdialysis recovery time and increased
mortality.\textsuperscript{7,8}

We performed a patient-blinded randomized
crossover study of patient-reported recovery time to
determine whether recovery time differs between HD
and HDF.

**METHODS**

**Study Design**

A patient-blinded, randomized, controlled, crossover design was
used. Patients were randomly assigned in a 1:1 ratio to receive 8
weeks of HD followed immediately by 8 weeks of online post-
dilution HDF, or vice versa. Patients were recruited from 2 satellite
dialysis units (Stobhill Hospital and Glasgow Royal Infirmary) and
consented by E.B., N.Z., J.R.S., or R.M. The study ran from July
2013 through March 2014. Randomization was conducted by E.B.
using a remote telephone-based system run by the Robertson
Centre for Biostatistics, University of Glasgow, United Kingdom.

The study was conducted in accordance with the ethics princi-
ples of the Declaration of Helsinki; all patients provided written
informed consent. Good Clinical Practice guidelines were fol-
lowed throughout. The West of Scotland Research and Ethics
Committee approved the study (13/WS/0010). Anonymized data
were sent to the Robertson Centre for Biostatistics for analysis.
Data analysis programs were developed prior to the release of
randomization codes to the study statistician.

**Patient Selection**

There were 198 patients screened for eligibility. Inclusion
criteria were receiving HD for more than 90 days, reliable vascular
access, and age of 18 years or older. Exclusion criteria were
currently receiving HDF (however, no recruited patients ended up
having previous HDF exposure), life expectancy less than 6
months, active neoplasia, recent (within 1 month) emergency
hospital admission, and unable to give informed consent or com-
plete questionnaires. Of 119 patients meeting these criteria, 100
underwent randomization, stratified by age (4 strata: 18-49, 50-59,
60-69, and ≥70 years) and sex (given sex differences noted in
recovery time reporting\textsuperscript{19}) into one of 2 groups: HD followed by
HDF, or HDF followed by HD. A separate randomization list was
generated for each stratum by a computer program, using the
method of randomized permuted blocks of length 4.

**Treatments**

Patients received 3 treatment sessions per week. Dialysis time,
dialyzer blood flow rate, dialysate flow rate, postdialysis weight,
and medications were kept constant unless changes were required
on clinical grounds. High-flux dialyzers (FX80 or FX100; Frese-
nius) were used to remain consistent with patients’ previous di-
alyzers and reduce the risk for inadvertent unblinding. Dialysate
composition was as follows: sodium, 138 mmol/L; potassium,
2 mmol/L; chloride, 108.5 mmol/L; bicarbonate, 32 mmol/L;
acetate, 3 mmol/L; calcium, 1.25 mmol/L; magnesium, 0.5 mmol/
L; and glucose, 1 g/L. Fresenius 5008 dialysis machines were
used. To further ensure patient blinding, dialysis machines were
turned away and the on-screen treatment modality notification was
covered. Dialysis unit staff were not blinded to treatment
allocation.

**Outcomes**

The primary outcome was reported recovery time in minutes. On
arrival for each session, patients were asked by the treating
nurse to state how long it took them to recover completely from
the preceding session. Secondary outcomes were frequency of
symptomatic hypotension; frequency of intradialytic clotting
events; pre-dialysis session serum concentrations of potassium,
phosphate, vitamin B\textsubscript{12}, parathyroid hormone, β\textsubscript{2}-microglobulin,
betaine, and interleukin 6; and Kt/V of urea.

**Measurement Methods**

Predialysis hematologic and biochemical tests were performed
following a 1-day treatment gap. Blood results from the middle
and end point of each treatment period were used in the analysis.
The nurses administering treatments were responsible for doc-
umenting primary and secondary end point data, as well as routine
dialysis session duties and data collection.

Although not prespecified outcomes, the frequency of other
adverse events (documented in free-text format by nursing staff),
change in quality-of-life scores (Kidney Disease Quality of Life–
Short Form [KDQOL-SF], version 1.3\textsuperscript{16}), change in dialysis dose,
and patients’ preferred dialysis modality at the end of the study
were also recorded. This was overseen by a dedicated research
nurse, who also administered the KDQOL-SF, version 1.3, ques-
tionnaires and processed blood samples.

**Sample Size Calculation**

This was based on pilot data demonstrating the variation in
recovery times in 100 patients over 3 consecutive HD sessions. To
detect a 20% absolute reduction in recovery time with 90% power,
we calculated that 82 patients would need to complete the study
(41 in each group). We planned to randomly assign 100 patients in
total, allowing for a dropout rate of 18%.

**Statistical Analyses**

Analyses were performed on an intention-to-treat basis. Baseline
characteristics and Charlson comorbidity scores\textsuperscript{20} were
compared between groups (HD then HDF vs HDF then HD) by
Fisher exact tests for categorical variables and $t$ tests for contin-
uous variables (or Wilcoxon tests if needed). Given that the dis-
tribution of recovery times was bimodal with a peak at zero,
recovery times were analyzed by crossover analysis with 2 mixed
models. A generalized linear mixed model (logistic regression
with random effect) was run to model the odds for a patient to
recover immediately (recovery time = 0 minutes) and a mixed
model was run to model the delayed (recovery time > 0 minutes)
recovery times (after a logarithmic transformation). The models
were combined by Monte Carlo simulations, then a parametric
bootstrap method was used to obtain the overall $P$ value.\textsuperscript{21}
Correlation between recovery times across successive sessions were
modeled with an autoregressive model of order 1 (in which each
recovery time is dependent on the previous session, and so on)
by treatment and time period. Additional analyses were run
for blood tests, dialysis data, Kt/V, and mean blood flow.
Hypotension, clotting, and adverse events were compared between
the treatments by relative risk (RR), and in order to take into ac-
count the crossover design, odds ratios (ORs) obtained by logistic
regression (with random effects) were also calculated. The
KDQOL survey was compared between treatments and baseline
with Friedman tests.
RESULTS

Study Participants

Table 1 shows baseline characteristics of 100 patients randomly assigned to receive HD then HDF or HDF then HD. Characteristics were generally representative of a typical prevalent HD cohort in the United Kingdom.²² Figure 1 gives an outline of the study flow.

Clinical Treatment Parameters

Treatment time and blood flow rate remained constant between HD and HDF (Table 2). Additionally, ultrafiltration volumes were similar and the mean convection volume for HDF treatments was 20.6 L. Pretreatment systolic blood pressure (SBP) was lower while participants were receiving HDF (143 vs 145 mm Hg; \( P = 0.03 \)); however, this difference was not demonstrated posttreatment (Table 2).

Primary Outcome: Posttreatment Recovery Time of HDF Versus HD

Recovery time data were available for 92% of all sessions. Recovery time for one-third of available sessions was reported as zero minutes (immediate), resulting in a bimodal distribution. To account for this, immediate and delayed (>0 minutes) recovery times were analyzed using separate models, then joined to obtain an overall \( P \) value, as described in the Statistical Analyses section in Methods. This demonstrated no overall difference in recovery time between HDF and HD (median values of 47.5 [IQR, 0-240] and 30 [IQR, 0-210] minutes, respectively; \( P = 0.9 \)). However, of interest, individual models for

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>All (N = 100)</th>
<th>HD Then HDF (n = 50)</th>
<th>HDF Then HD (n = 50)</th>
<th>( P^a )</th>
</tr>
</thead>
<tbody>
<tr>
<td>Female sex</td>
<td>39 (39)</td>
<td>20 (40)</td>
<td>19 (38)</td>
<td>0.9</td>
</tr>
<tr>
<td>Age, y</td>
<td>65 ± 14</td>
<td>65 ± 15</td>
<td>66 ± 13</td>
<td>0.6</td>
</tr>
<tr>
<td>Access</td>
<td></td>
<td></td>
<td></td>
<td>0.2</td>
</tr>
<tr>
<td>Fistula</td>
<td>68 (68)</td>
<td>31 (62)</td>
<td>37 (74)</td>
<td></td>
</tr>
<tr>
<td>Graft</td>
<td>1 (1)</td>
<td>0 (0)</td>
<td>1 (2)</td>
<td></td>
</tr>
<tr>
<td>Central venous catheter</td>
<td>31 (31)</td>
<td>19 (38)</td>
<td>12 (24)</td>
<td></td>
</tr>
<tr>
<td>Antihypertensive medication use</td>
<td>66 (66)</td>
<td>31 (62)</td>
<td>35 (70)</td>
<td>0.5</td>
</tr>
<tr>
<td>SBP, mm Hg</td>
<td>143 ± 20</td>
<td>144 ± 22</td>
<td>142 ± 18</td>
<td>0.7</td>
</tr>
<tr>
<td>DBP, mm Hg</td>
<td>69 ± 12</td>
<td>67 ± 12</td>
<td>70 ± 11</td>
<td>0.3</td>
</tr>
<tr>
<td>Post-HD weight, kg</td>
<td>74 [62-87]</td>
<td>76 [64-88]</td>
<td>71 [61-86]</td>
<td>0.2⁶</td>
</tr>
<tr>
<td>UF volume, mL</td>
<td>1,819 ± 694</td>
<td>1,995 ± 681</td>
<td>1,644 ± 668</td>
<td>0.01</td>
</tr>
<tr>
<td>Primary renal diagnosis</td>
<td></td>
<td></td>
<td></td>
<td>0.7</td>
</tr>
<tr>
<td>Glomerular</td>
<td>26 (26)</td>
<td>13 (26)</td>
<td>13 (26)</td>
<td></td>
</tr>
<tr>
<td>Tubulointerstitial</td>
<td>11 (11)</td>
<td>5 (10)</td>
<td>6 (12)</td>
<td></td>
</tr>
<tr>
<td>Systemic</td>
<td>30 (30)</td>
<td>17 (34)</td>
<td>13 (26)</td>
<td></td>
</tr>
<tr>
<td>Hereditary</td>
<td>10 (10)</td>
<td>3 (6)</td>
<td>7 (14)</td>
<td></td>
</tr>
<tr>
<td>Miscellaneous</td>
<td>23 (23)</td>
<td>12 (24)</td>
<td>11 (22)</td>
<td></td>
</tr>
<tr>
<td>Waitlisted for transplant</td>
<td>21 (21)</td>
<td>13 (26)</td>
<td>8 (16)</td>
<td>0.3</td>
</tr>
<tr>
<td>White</td>
<td>99 (99)</td>
<td>49 (98)</td>
<td>50 (100)</td>
<td>0.5</td>
</tr>
<tr>
<td>Smoking status</td>
<td></td>
<td></td>
<td></td>
<td>0.5</td>
</tr>
<tr>
<td>Current smoker</td>
<td>27 (27)</td>
<td>11 (22)</td>
<td>16 (32)</td>
<td></td>
</tr>
<tr>
<td>Ex-smoker</td>
<td>30 (30)</td>
<td>17 (34)</td>
<td>13 (26)</td>
<td></td>
</tr>
<tr>
<td>Never smoked</td>
<td>43 (43)</td>
<td>22 (44)</td>
<td>21 (42)</td>
<td></td>
</tr>
<tr>
<td>Diabetes</td>
<td>26 (26)</td>
<td>15 (30)</td>
<td>11 (22)</td>
<td>0.5</td>
</tr>
<tr>
<td>Ischemic heart disease</td>
<td>37 (37)</td>
<td>20 (40)</td>
<td>17 (34)</td>
<td>0.7</td>
</tr>
<tr>
<td>Peripheral vascular disease</td>
<td>19 (19)</td>
<td>14 (28)</td>
<td>5 (10)</td>
<td>0.04</td>
</tr>
<tr>
<td>Stroke</td>
<td>17 (17)</td>
<td>10 (20)</td>
<td>7 (14)</td>
<td>0.6</td>
</tr>
<tr>
<td>History of neoplasia</td>
<td>7 (7)</td>
<td>6 (12)</td>
<td>1 (2)</td>
<td>0.1</td>
</tr>
<tr>
<td>Charlson comorbidity score</td>
<td>7 ± 2</td>
<td>7 ± 2</td>
<td>6 ± 2</td>
<td>0.4⁵</td>
</tr>
</tbody>
</table>

Note: Values for categorical variables are given as count (percentage); values for continuous variables, as mean ± standard deviation or, in the case of non-normally distributed data, median [interquartile range].

Abbreviations: DBP, diastolic blood pressure; HD, high-flux hemodialysis; HDF, hemodiafiltration; SBP, systolic blood pressure; UF, ultrafiltration.

²²P values derived from t test for continuous variable (or Wilcoxon) and Fisher exact test for categorical variables; 100% of data is reported for all variables.

³Mean of 3 months of dialysis data prior to randomization.
immediate and delayed recovery time showed that patients were more likely to report immediate recovery while receiving HDF treatment (34.4% for HDF vs 32.2% for HD; OR, 1.37; 95% CI, 1.08-1.74; \( P = 0.01 \); Fig 2A). Conversely, HDF resulted in longer delayed recovery times (for HDF vs HD, medians of 150 [IQR, 60-420] vs 137 [IQR, 64-296] minutes, geometric means of 120 [IQR, 30-270] vs 104 [IQR, 64-296] minutes, respectively; \( P , 0.001 \); Fig 2B).

Table S1 (provided as online supplementary material) shows baseline characteristics comparing patients who reported at least one immediate recovery time with those who reported only delayed recovery times. There was no interaction between treatments and the order in which they were given (treatment effect) for immediate (\( P = 0.8 \)) or delayed (\( P = 0.6 \)) recovery times. Although there was no evidence of a difference in recovery time reporting between the first and second period of the study (time effect) for delayed recovery times (OR for delayed recovery, second vs first study period, 0.99; 95% CI, 0.89-1.09; \( P = 0.8 \)), patients were more likely to report immediate recovery during the first half of the study (OR for immediate recovery, second vs first study period, 0.77; 95% CI, 0.61-0.98; \( P = 0.04 \)), regardless of treatment (Tables S2 and S3).

A washout period was not possible given the nature of the treatments. However, with regard to carryover effect, the correlation between sessions decreased quickly and was \(<5\%\) for both treatments after 9 sessions. To test for the influence of carryover effect, analyses were repeated by removing up to and including the first 9 sessions of each treatment. Immediate, delayed, and overall recovery times remained equivalent after these alterations (Table S4).

Secondary Outcomes

**Symptomatic Hypotension**

HDF was associated with an increased rate of symptomatic hypotension compared to HD (8.0% vs 5.3%; RR, 1.52; 95% CI, 1.2-1.9; \( P < 0.001 \); Table 3). Antihypertensive dosing was increased in 3 patients while on HD therapy and 1 patient while on HDF therapy (\( P = 0.6 \)); similarly, dosing was reduced in 3 patients while on HD therapy and 1 patient while on HDF therapy (\( P = 0.6 \)).

**Clotting Risk**

The intradialytic tendency to clotting was higher during HDF than HD (1.8% vs 0.7%; RR, 2.7; 95% CI, 1.5-5.0; \( P = 0.002 \); Table 3). There was no difference in tinzaparin dosing (2,773 ± 147 and 2,740 ± 140 units for HDF and HD, respectively; \( P = 0.9 \)).

**Laboratory Values**

None of the prespecified laboratory measurements showed statistically significant differences between HDF and HD treatments (Table 4). Blood samples were stored for additional analyses (eg. \( \beta \)-2-microglobulin,

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**Table 2.** Delivered Treatment and Peritreatment Blood Pressure: HD Versus HDF

<table>
<thead>
<tr>
<th>Variable</th>
<th>HD</th>
<th>HDF</th>
<th>Crossover P</th>
<th>Data Available, %</th>
</tr>
</thead>
<tbody>
<tr>
<td>Session length, min</td>
<td>250 ± 17</td>
<td>250 ± 17</td>
<td>0.1</td>
<td>94</td>
</tr>
<tr>
<td>Blood flow rate, mL/min</td>
<td>315 ± 27</td>
<td>313 ± 28</td>
<td>0.6</td>
<td>93</td>
</tr>
<tr>
<td>Ultrafiltration volume, mL</td>
<td>1,749 ± 718</td>
<td>1,723 ± 672</td>
<td>0.4</td>
<td>94</td>
</tr>
<tr>
<td>Convection volume, L</td>
<td>NA</td>
<td>20.6 ± 4.6</td>
<td>NA</td>
<td>82</td>
</tr>
<tr>
<td>Pretreatment SBP, mm Hg</td>
<td>145 ± 21</td>
<td>143 ± 21</td>
<td>0.03</td>
<td>93</td>
</tr>
<tr>
<td>Pretreatment DBP, mm Hg</td>
<td>69 ± 12</td>
<td>69 ± 12</td>
<td>0.1</td>
<td>93</td>
</tr>
<tr>
<td>Posttreatment SBP, mm Hg</td>
<td>126 ± 20</td>
<td>125 ± 18</td>
<td>0.08</td>
<td>93</td>
</tr>
<tr>
<td>Posttreatment DBP, mm Hg</td>
<td>64 ± 10</td>
<td>63 ± 10</td>
<td>0.1</td>
<td>93</td>
</tr>
</tbody>
</table>

*Note: Unless otherwise indicated, values are given as mean ± standard deviation.*

**Abbreviations:** DBP, diastolic blood pressure; HD, high-flux hemodialysis; HDF, hemodiafiltration; NA, not applicable; SBP, systolic blood pressure.
betaine, interleukin 6) but these samples have not been processed.

Other Measures

Other Laboratory Measures

As shown in Table 4, there were small but statistically significant differences in serum albumin (3.2 vs 3.3 g/dL for HDF and HD, respectively; \( P, 0.001 \)) and chloride levels (101 vs 100 mEq/L for HDF and HD; \( P = 0.02 \)).

Other Adverse Events

Other adverse events were grouped according to those that may or may not have been related to blood pressure (BP) changes or fluid shifts (Table 3). Although the frequency of events not typically associated with BP or fluid shifts was similar between groups, there were more instances of events more likely to be related to BP or fluid shifts during HDF treatment compared to HD (5.3% vs 3.0%; RR, 1.8; 95% CI, 1.3-2.5; \( P < 0.001 \)).

Quality-of-Life Scores

Patients completed KDQOL-SF, version 1.3, questionnaires\(^{19} \) during the study. Patients scored physical health lower (33 ± 10) than mental health (44 ± 10) at baseline. There was no difference in physical health composite scores (33 ± 9 for both HDF and HD; \( P = 0.9 \)) or mental health composite scores (44 ± 11 vs 43 ± 12 for HDF and HD, respectively; \( P = 0.6 \)) after 8 weeks of each treatment.

Withdrawals/Deaths

There were 9 withdrawals during the study, 2 occurring after randomization but before starting the

---

table 3

<table>
<thead>
<tr>
<th>Variable</th>
<th>HD Sessions</th>
<th>HDF Sessions</th>
<th>RR (95% CI)</th>
<th>( P )</th>
<th>Data Available, %</th>
</tr>
</thead>
<tbody>
<tr>
<td>Symptomatic hypotension(^{a})</td>
<td>112 (5.2)</td>
<td>168 (8)</td>
<td>1.52 (1.21-1.92)</td>
<td>&lt;0.001</td>
<td>97</td>
</tr>
<tr>
<td>AEs potentially related to BP/fluid shifts(^{b})</td>
<td>61 (3.0)</td>
<td>109 (5.3)</td>
<td>1.81 (1.33-2.46)</td>
<td>&lt;0.001</td>
<td>93</td>
</tr>
<tr>
<td>AEs not classically related to BP/fluid shifts(^{c})</td>
<td>88 (4.3)</td>
<td>87 (4.3)</td>
<td>1.00 (0.75-1.34)</td>
<td>0.9</td>
<td>93</td>
</tr>
<tr>
<td>Extra tinzaparin dose(s) or clotting of circuit(^{d})</td>
<td>14 (0.7)</td>
<td>37 (1.8)</td>
<td>2.68 (1.46-5.00)</td>
<td>0.002</td>
<td>97</td>
</tr>
</tbody>
</table>

Note: Unless otherwise indicated, values are given as number of events (percentage). Multiple episodes within 1 session were treated as a single event. Odds ratios taking into account the crossover design were also calculated and were almost identical to RRs.

Abbreviations: AE, adverse event; BP, blood pressure; CI, confidence interval; HD, high-flux hemodialysis; HDF, hemodiafiltration; RR, relative risk.

\(^{a}\)Defined as a decrease in systolic BP \( \geq 20 \) mm Hg requiring reduction or cessation of ultrafiltration and/or need for intravenous fluid bolus or head-down tilt of dialysis chair.

\(^{b}\)Breathlessness, cramp (normal BP), dizzy/lightheaded, fall, headache, venous pressures erratic, clotted needle, or restless legs.

\(^{c}\)Aches in bones, arm pain, back pain, bleeding, constipation, diarrhea, feeling cold, feeling down, feeling hot, generally unwell, heavy legs, increased lethargy, infection (given antibiotics), itch, leg pain, nausea, stomach pains, sweating, swollen abdomen, and vomiting.

\(^{d}\)Defined as either an increase in venous pressure requiring additional anticoagulant dosing or clotting of the extracorporeal circuit.
trial as a result of emergency hospital admissions. Four patients died during the study. Further details are given in Table S5.

**Maintenance of Blinding and Patient Preferences**

Blinding to treatment was maintained in 84 of the 86 patients who completed the study. While still blinded, these 84 patients were asked which treatment they preferred; 52 had no preference, 21 preferred HDF, and 11 preferred HD. Eight patients felt able to guess the treatment order based on their symptoms; 5 were correct and 3 were incorrect.

**DISCUSSION**

This patient-blinded, randomized, controlled, crossover trial showed no difference in the primary outcome of patient-reported posttreatment recovery time between HDF and HD (medians of 47.5 [IQR, 0-240] and 30 [IQR, 0-210] minutes, respectively; \( P = 0.9 \)). However, recovery time data had a bimodal distribution, with more than one-third of recovery times being recorded as zero minutes (immediate). Therefore, analysis of immediate recovery was performed separately from delayed recovery, with interesting yet disparate results. Treatment with HDF was associated with a significantly higher chance of immediate recovery \( (P = 0.01) \), but resulted in significantly longer recovery time than HD for those who had delayed recovery \( (P < 0.001) \).

We hypothesize that inherent differences among patients reporting immediate and delayed recovery times may have influenced treatment tolerance. For example, patients with fewer comorbid conditions may report more symptomatic benefit from enhanced small- and middle-molecule clearance attributed to convective treatments, whereas patients with more comorbid conditions may tolerate HDF less well. A recent subanalysis of the DOPPS cohort showed that recovery time, based on a categorical measure at a single time, was similar in patients on HDF and HD therapy \( (OR, 1.08; 95\%\; CI, 0.87-1.35) \). Interestingly, longer recovery time was associated with increased mortality, as well as adverse clinical features such as older age, higher body mass index, and diabetes. We found that patients who reported immediate recovery were significantly more likely at baseline to be lighter \( (73\; vs\; 86\; kg;\; P = 0.008) \) and active on the transplant waiting list \( (29\% \; vs\; 6\%;\; P = 0.009) \) than patients who reported only delayed recovery times, but other baseline characteristics were similar (Table S1).

Patient-reported posttreatment recovery time is inherently subjective, difficult to measure robustly, and at risk for reporting or recall bias. Our choice of crossover design was made in the face of these potential issues in order to abrogate the influence of interpatient variability. The question “How long does it take you to recover from a dialysis session?” has been validated as easy to understand and stable in the face of retest, yet sensitive to alterations in treatment, and this was the same question administered to the DOPPS cohort. The former study asked the
question at 3-month intervals (recording the answer as a continuous variable), and in the latter study, the question was asked once at the beginning of the study and collected as a categorical variable. In order to achieve sufficient granularity, we recorded recovery time as a continuous variable; how this compares to a categorical approach is not known. In addition, we adapted the question asked of the study patients to “How long did it take you to recover from your previous treatment session” given that this was being asked at every session. The differences noted between treatment modalities in immediate and delayed recovery time reporting in this study suggest that this approach was sufficiently sensitive to detect changes in recovery time.

We found no difference in HRQoL scores between HD and HDF, which is consistent with the result of the overall primary outcome. Longer postdialysis recovery time has been associated with poorer HRQoL scores. Additionally, associations have also been drawn between lower HRQoL scores and patient mortality.

There was an increase in risk for intradialytic clotting with HDF, which in the majority of cases was mitigated by an increase in anticoagulant dose prior to clotting of the circuit. Clotting resulting in the need to change the circuit was infrequent, occurring in 0.62% of HDF sessions and 0.38% of HD sessions. Postdilution HDF provides more efficient convective clearance compared with predilution HDF, but has previously been associated with an increase in clotting events compared to HD. There was no difference in mean anticoagulant doses between treatments in our study, perhaps due to the small number of clotting events prompting an increase in subsequent anticoagulant doses.

Unexpectedly, we found that HDF was associated with increased symptomatic hypotension. Eighty percent of these episodes were classed as mild, requiring only cessation of ultrafiltration. There was also a small reduction in pretreatment SBP in those receiving HDF, whereas posttreatment SBP was equivalent between treatments. Convective treatments are often considered to have a hemodynamic advantage over HD. This has been described in studies comparing low-flux HD to hemofiltration or HDF, but not in a study in which both modalities were performed under matched conditions, and we are not aware of blinded trials analyzing this outcome. When comparing HD to HDF, Ok et al found no difference in symptomatic hypotension (defined as a decrease in SBP > 30 mm Hg requiring saline solution infusion), and in the ESHOL (Estudio de Supervivencia de Hemodiafiltración On-Line) study, the incidence of symptomatic hypotension (not clearly defined) was lower in the HDF group (rate ratio, 0.72; 95% CI, 0.68-0.77; P < 0.001). The reasons behind the disparate findings are unclear, but importantly, in our study, there were no significant changes in antihypertensive treatment, ultrafiltration volume, blood flow, or dialysis dose between treatments.

With regard to blood parameters, it is worth highlighting the small reduction in serum albumin levels during HDF treatment compared to HD (3.2 vs 3.3 g/dL, respectively; P < 0.001). Data for serum albumin levels in previous studies comparing HDF to HD are variable. Ok et al found that albumin levels were lower in the HDF group, whereas in ESHOL, there was no difference between treatments. However, serum albumin levels decreased in both groups compared to baseline over the course of both these randomized studies. Even small reductions in serum albumin levels are strongly associated with reduced survival in HD patients. Therefore, it is important to be aware of such changes, particularly in the presence of pre-existing hypoalbuminemia.

In light of these findings, it is conceivable that more frequent episodes of hypotension and lower serum albumin levels may have influenced recovery time adversely in the HDF group. However, we found no difference in results of the primary outcome when adjustments were made for BP and serum albumin level (Table S4).

The success of patient blinding exceeded our expectations and is worthy of note; this should be considered a realistic strategy for future studies investigating extracorporeal treatments. Due to the practicalities of administering these treatments, double-blinding was not possible, and the treating nurses were responsible for collecting end point data from patients. However, statistical analyses were performed on anonymized allocation-blinded data. Inclusion criteria were designed to be as broad as possible to serve as a representative cohort of long-term HD patients. It may be that this in part limited our ability to achieve HDF convection volumes described in previous studies. For example, in our study, the arteriovenous fistula rate (68%) was lower than in the ESHOL online-HDF group (89%), as were mean blood flow rates (313 mL/min in our study vs 389 mL/min after 36 months in the ESHOL online-HDF group). These factors, in addition to treatment time, are known to be important in determining achieved convection volumes. In our study, mean treatment time was 250 minutes, longer than the 237 minutes reported after 36 months in the ESHOL online-HDF group, suggesting that blood flow rate was one of the most important factors limiting achieved convection volume.

The single-center (although multiple satellite unit) design is a limitation. Additionally, the predominance
of patients of European ancestry, although representative of our case-mix, will be less applicable in other regions. It is also possible that the nursing staff’s longer term experience with HD compared to HDF could have had an influence on results. However, this situation is likely to be the case in most nephrology units, and crossover analyses demonstrated more immediate recovery times for both treatments during the first treatment period, with no difference in delayed recovery times between treatment periods, suggesting that increasing staff experience in delivering HDF treatment is likely to have had minimal influence.

To our knowledge, this is the first RCT to examine recovery time in the setting of HD and HDF, and we found no overall difference between treatments. However, HDF was associated with both more episodes of immediate recovery and longer recovery time if recovery was not immediate. Although more recent data may suggest a beneficial role for online HDF with high convection volumes, many patients may be unable to achieve such convection volumes and therefore neither treatment is definitively superior for all patients.\(^{32}\) The overall result of our primary outcome is consistent with the findings of previous observational studies,\(^{18,24,25}\) however, it is difficult to infer which of the many variables associated with extracorporeal treatments may be responsible for effects on recovery time or mortality. For example, more frequent (6 times per week) HD has been associated with a reduction in recovery time compared to thrice-weekly HD,\(^{33}\) as well as improvements in areas such as BP and serum phosphate control, left ventricular mass, and HRQoL.\(^{34-36}\) Whether this leads to improvements in patient survival remains a point of debate.\(^{37-40}\) With this in mind and in light of the association between recovery time and mortality as identified in the DOPPS cohort,\(^{18}\) a large-scale, blinded, multicenter RCT of HDF and HD powered to examine morbidity and mortality outcomes while taking into account recovery time reporting and quality-of-life scores across a representative patient group would be of benefit.

Debate remains regarding the clinical case for HDF over HD. Patient preference and shared decision making are increasingly prioritized in clinical practice, and these data may further inform the discussion around choice of extracorporeal treatments.

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**Peer Review:** Evaluated by 2 external peer reviewers, a Statistical Editor, a Co-Editor, and Editor-in-Chief Levey.

**SUPPLEMENTARY MATERIAL**

Table S1: Baseline characteristics of participants reporting ≥1 immediate recovery time vs none.

Table S2: Percentage of immediate recovery times according to treatment sequence.

Table S3: Delayed recovery times according to treatment sequence.

Table S4: Sensitivity analysis of recovery times: comparison of delayed and immediate times (HDF vs HD).

Table S5: Withdrawals and deaths during study period.

Note: The supplementary material accompanying this article (http://dx.doi.org/10.1053/j.ajkd.2016.10.025) is available at www.ajkd.org


