Theory of Mind and Psychosocial Characteristics in Older Men

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Abstract

The extent to which early-life cognitive ability shapes individuals’ social functioning throughout life, in the context of later-life factors, is unknown. We investigated performance on the Faux Pas test (FP) in relation to psychosocial characteristics and childhood intelligence scores in 90 healthy older men. FP performance was associated with close social network size but not social contact, social support, or loneliness when accounting for both childhood and later-life intelligence, affect, personality, and socio-demography. We add to a growing literature on associations between ToM and intelligence, affect, and personality.

Keywords: Theory of mind, Faux Pas test, psychosocial characteristics, individual differences, older age

Abstract = 88 words

Manuscript = 3,500 words
Introduction

Socioemotional selectivity theory (Carstensen, 2006) posits that older individuals strategically prune their social networks in favor of social environments that are more emotionally satisfying (English & Carstensen, 2014). The positive emotions experienced by older adults during social interactions are thought to improve their emotional well-being (Scheibe & Carstensen, 2010) and social functioning (Luong, Charles, & Fingerman, 2011; see Charles & Carstensen, 2009).

The formation and maintenance of successful social relationships in older adults may also rely on Theory of mind (ToM). ToM is the ability to ascribe mental states to oneself and others to explain and predict behavior (Premack & Woodruff, 1978). Deficits in performance-based ToM correspond to impairments in real-world social functioning among clinical groups (Bishop-Fitzpatrick, Mazefsky, Eack, & Minshew, 2017) and typically developing children (Caputi, Lecce, Pagnin, & Banerjee, 2012). However, while ToM declines in older age (Henry, Phillips, Ruffman, & Bailey, 2013), it remains poorly understood the role that ToM plays in older adults’ social lives.

In one of the few studies exploring the relationship between ToM and psychosocial functioning in later adulthood, Bailey, Henry, and von Hippel (2008) demonstrated that reduced mental state understanding was associated with decreased social participation in older adults. Yeh (2013) later showed that ToM performance was related to self-reported social skills in older age. Also, Lecce et al. (2017) found links between older adults’ ToM performance and self-perceived friendship indicators. Nonetheless, not only does the robustness of these associations require further investigation, but also whether they are potentially modulated by individual differences in non-social psychological domains.
The correspondence between ToM and psychosocial characteristics might be partly affected by cognitive variation. The two-systems account of ToM proposes that mental state understanding relies on both the readily triggered implicit processes of sensory perception (e.g., face processing) and the consciously guided explicit processes of higher cognitive skills such as language and reasoning (Frith & Frith, 2008; Apperly & Butterfill, 2009). Indeed, ToM has been related to verbal ability and abstract reasoning in young adults (Ahmed & Miller, 2011), and abstract reasoning in older adults (Cox et al., 2014). Meinhardt-Injac, Daum, Meinhardt, and Persike (2018) found that a verbal latent factor, but not an abstract reasoning one, predicted ToM abilities in young adults. To our knowledge, however, the independent influences of language and abstract reasoning on associations between ToM and psychosocial characteristics have not been examined in older age.

Furthermore, while certain cognitive functions (including fluid intelligence) deteriorate, on average, with age (Singh-Manoux et al., 2012), individual differences in intelligence remain relatively stable from childhood to later adulthood (Deary, 2014). While such measures of general cognitive ability are related to ToM in children (Ibanez et al., 2013) and older adults (Cox et al., 2014; this sample), the lifelong stability of intelligence poses questions about how much later-life ToM and social support might simply reflect a lifelong association between early life intelligence and social abilities. However, this possibility, in the context of other important late-life factors, has yet to be investigated.

Links between ToM and psychosocial factors may also be influenced by affect and personality. Major depressive disorder (MDD; Bora & Berk, 2016) and social anxiety disorder (Washburn, Wilson, Roes, Rnic, & Harkness, 2016) have been coupled with ToM impairments in young and/or older adulthood. Regarding personality, agreeableness, for example, has been shown to correlate with ToM performance in young adults (alongside neuroticism; Nettle & Liddle, 2008) and moderate day-to-day interpersonal conflicts in
adolescents (Jensen-Campbell & Graziano, 2001). Yet, the effects of affect and personality on links between ToM and psychosocial functioning remain seemingly unexplored in older age.

The Faux Pas test (FP; Stone, Baron-Cohen, & Knight, 1998) is a widely-used task that is administered to assess age-related decline in ToM (Henry et al., 2013). Participants are presented with twenty written stories, ten of which contain a faux pas. Participants should identify whether the protagonists’ spoken words may have unintentionally upset other characters in the stories; this requires an understanding of others’ cognitive and affective states.

The present study aimed to characterize relationships between ToM, as measured by the FP test, and psychosocial characteristics in older age, where impaired mental state understanding putatively entails decreased social participation. The second goal was to determine whether early-life cognitive ability influences the relationship between ToM and psychosocial characteristics in later life or whether there are additional influences from later-life non-social factors, such as crystallized ($g_c$) and fluid ($g_f$) intelligence, depressive and anxious symptoms, and personality traits, but also socio-demography.

**Method**

**Participants**

Ninety males aged 73.2 to 75.0 years from the Lothian Birth Cohort 1936 (LBC1936) participated. Cohort members were born in 1936 and tested at age ~11 on the Moray House Test (MHT) for the Scottish Mental Survey 1947 (Deary, Whalley, & Starr, 2009; Scottish Council for Research in Education, 1949). Raw MHT scores (maximum = 76) were converted into IQ scores (M = 100, SD = 15) and corrected for age (in days) at testing. Participants were recruited from Edinburgh and surrounding areas for follow-up testing at age ~70 years (Wave 2).
1; Deary et al., 2007), and returned three years later for further testing (Wave 2; Deary, Gow, Pattie, & Starr, 2012).

The current sample was invited to participate in a cortisol study (Cox et al., 2014; 2015a,b; 2017). The inclusion criteria included: male (due to sex differences in glucocorticoid secretion patterns); ≥ 24 on the Mini-Mental State Examination (MMSE; Folstein, Folstein, & McHugh, 1975); < 11 on the depression subscale of the Hospital Anxiety and Depression Scale (HADS; Zig mond & Snaith, 1983); a structural brain magnetic resonance imaging scan within 1.5 years; and no current antidepressant or glucocorticoid medication use. Ethical permission was obtained from the Multi-Centre Research Ethics Committee for Scotland (MREC/01/0/56), the Lothian Research Ethics Committee (LREC/2003/2/29), and the Scotland A Research Ethics Committee (07/MRE00/58). All volunteers gave written informed consent.

At Wave 2, participants indicated their relationship status, dichotomously coded as single/divorced/widowed (0) versus married/cohabiting/other (1), and living arrangement, dichotomously coded as living alone (0) versus living with others (1). An index of local deprivation (Scottish Index of Multiple Deprivation; SIMD; Scottish Executive, 2004) was derived for each participant at Wave 1, assessing employment, income, health, education, access, crime, and housing. SIMD (henceforth, reported as local deprivation) was measured on a ratio scale, with a score of 1 indicating the most and 6505 the least deprived data zone.

Measures

The Faux Pas test. Wave 2 participants undertook the FP test individually. Ten of the twenty stories contained a verbal faux pas. Participants read each story at their own pace and indicated when they were finished. They were then read a series of eight questions while the story remained in front of them. The first assessed whether participants recognized that a faux
pas had been committed. The next four assessed cognitive ToM: who committed the faux pas; the inappropriateness of what was said; the speaker’s motivations; and a story character’s beliefs. The final question assessed empathy/affective ToM: how the characters felt. Finally, two control questions assessed whether participants understood the story. The main outcome variable was the total number of correct responses for the ten faux pas stories, excluding responses to the empathy/affective ToM and control questions. Each correct response was awarded 1 point (maximum = 50). Scoring was according to the Stone et al. (1998) guidelines. The FP test is considered a valid measure of ToM (Baron-Cohen, O’Riordan, Stone, Jones, & Plaisted, 1999), and has good internal consistency (Cronbach’s α = .91) and test-retest reliability (Cronbach’s α = .89; Yeh, Hua, & Liu, 2009).

Social contact. Wave 1 participants completed eight yes/no items relating to social life (Gow, Corley, Starr, & Deary, 2013). The four questions regarding social contact (Cronbach’s α = .67) – such as: “In the last 2 weeks, excluding people you live with, have you seen a friend to have a chat to?” – were summed (for all questions, see Supplemental Table 1). Higher scores indicated greater social contact.

Social support. Wave 1 participants completed a twelve-item questionnaire adapted from the Social Support Questionnaire (Short Form; SSQ; Sarason, Sarason, Shearin, & Pierce, 1987). It comprised six questions regarding the level of social support received, measured on a five-point scale from None of the time to All of the time, such as: “How often could you count on people to console you when you were very upset?” (for all questions, see Supplemental Table 2). Higher scores indicated greater social support.

Loneliness. Wave 2 participants were asked from the European Social Survey (2006): “How often have you felt lonely during the past week?” The item was measured on a four-point scale ranging from None or almost none of the time to All or almost all of the time. Since 74 out of 90 participants (82.2%) reported feeling lonely none or almost none of the
time, the data were collapsed into a dichotomous variable coded as not lonely (0) versus lonely (1).

**Close social network size.** Wave 1 participants were asked to estimate the size of their close social networks in response to: “About how many ‘close’ friends and ‘close’ relatives do you have (‘close’ meaning people that you feel at ease with, could talk to about what was on your mind, and could call on for help)?”

**Crystallized intelligence (g<sub>c</sub>).** To measure g<sub>c</sub>, the Wechsler Test of Adult Reading (WTAR; Wechsler, 2001) was administered at Wave 2. It requires the pronunciation of 50 irregularly spelled words; 1 point for each correct pronunciation was awarded, and a score out of 50 was used as the outcome variable.

**Fluid intelligence (g<sub>f</sub>).** G<sub>f</sub> was assessed at Wave 2 using five subtests from the Wechsler Adult Intelligence Scale-III UK (WAIS-III UK; Wechsler, 1997a): Block Design, Letter-Number Sequencing, Matrix Reasoning, Digit Symbol, and Symbol Search; and one subtest from the Wechsler Memory Scale-III UK (WMS-III UK; Wechsler, 1997b): Backward Digit Span. In all subtests, higher scores indicated higher g<sub>f</sub>.

**Depression and anxiety.** The HADS (Zigmond & Snaith, 1983) was completed at Wave 2, with seven items examining symptoms of depression and seven of anxiety. Both subscales had a maximum score of 21; higher values represented more depressive and anxious symptoms.

**Personality.** The International Personality Item Pool (IPIP; Goldberg, 1999) was administered at Wave 2 to assess Big-Five personality traits (Costa & McCrae, 1992): extraversion; agreeableness; conscientiousness; emotional stability (reverse of neuroticism); and intellect (henceforth, reported as openness). Higher scores indicated greater personality trait presence, and each domain allowed for a maximum score of 40.

See Supplemental Table 3 for the assessment wave for each study variable.
Statistical analyses

Statistical analyses were performed in R version 3.2.2 (“Fire Safety”; R Core Team, 2016). Social support was derived by entering the six SSQ questions (Cronbach’s $\alpha = .94$) into a Principal Component Analysis (PCA) using the “principal” function from the “psych” package, using data from all LBC1936 Wave 2 participants ($N = 1091$), from which the scores for the FP subsample ($N = 90$) were extracted. The first unrotated principal component accounted for 76% of the variance (all loadings ≥ .84; Supplemental Table 2).

To derive a general measure of $g_c$, a PCA was conducted using five WAIS-III UK (Block Design, Letter-Number Sequencing, Matrix Reasoning, Digit-Symbol, and Symbol Search) and one WMS-III UK subtests (Backward Digit Span), again using all available data from the LBC1936 at Wave 2 and then extracting the scores for the FP subsample. The first unrotated principal component accounted for 51% of the variance (Cronbach’s $\alpha = .73$; all loadings > .65; Supplemental Table 4).

We used Welch’s t-tests (for ratio variables) and Pearson’s chi-squared tests (for categorical and ordinal variables) to examine whether those who performed the FP test ($N = 90$) differed significantly from those who did not ($N = 1001$) on psychosocial, cognitive, affective, personality, and socio-demographic data. Next, we examined bivariate associations between all study variables (Pearson’s $r$).

We then used multiple linear regression – simultaneously modelling intelligence, affect, personality, and socio-demography – to ascertain the unique contributions each variable made to FP performance. Living arrangement was dropped from the model given its high variance inflation alongside relationship status. Then, we regressed MHT scores onto $g_c$ and $g_f$, and ran the model again with these residuals, thereby accounting for childhood intelligence. Finally, associations between FP performance and our psychosocial outcomes, treated as dependent variables, were explored using hierarchical regression. Separate models
were fitted for each psychosocial variable to examine the variance accounted for by FP performance alongside other potentially confounding measures. With each model iteration, predictors were added incrementally: 1) FP; 2) age (in days); 3) intelligence; 4) symptoms of depression and anxiety; 5) five personality traits; 6) relationship status and local deprivation (again, living arrangement was not included). We used multiple linear regression except for loneliness, where we used binomial logistic regression. P-values were corrected for multiple comparisons using False Discovery Rate (FDR; Benjamini & Hochberg, 1995) with “p.adjust” from the “stats” package. Variance inflation was ascertained for each model fit using “vif” from the “car” package. Finally, we regressed MHT scores onto gc and gf, and ran the models again using these residuals to take account of childhood intelligence.

**Results**

Descriptive statistics and tests of differences between the FP subsample and the LBC1936 (excluding the FP subsample) are reported in Supplemental Table 5.

Bivariate associations between the study variables within the FP subsample are shown in Table 1. They revealed that better FP performance correlated with higher MHT scores ($r = .591, p < .001$), higher gc ($r = .588, p < .001$), higher gf ($r = .482, p < .001$), higher openness ($r = .311, p = .003$) and emotional stability ($r = .222, p = .036$), and fewer depressive symptoms ($r = -.278, p = .008$). Better FP performance also correlated with smaller close social network size ($r = -.397, p < .001$), but not with any other psychosocial characteristic ($rs \leq |.199|, ps \geq .060$).

- Insert Table 1 around here -
Using multiple regression, only $g_c$ ($\beta = .367, p < .001$), $g_f$ ($\beta = .216, p = .026$), and depressive symptoms ($\beta = -.227, p = .025$) made significant unique contributions to FP performance, with the overall model accounting for 39% of the total variance. $G_c$ was the only variable to survive FDR correction ($q < .007$); see Supplemental Table 6. When considering childhood intelligence, $g_c$ and $g_f$ no longer significantly contributed to FP performance ($p \geq .09$). Only depressive symptoms ($\beta = -.277, p = .030$) were a significant predictor but this did not survive FDR correction (Supplemental Table 7).

Finally, we determined whether FP performance predicted psychosocial characteristics (Table 2). Higher FP performance was significantly associated with smaller close social network size in all model iterations, and this relationship survived correction for multiple testing ($\beta$ range = $-.501$ to $-.435$, $p$-value range = $< .001$ to $.010$). Better FP performance was also significantly related to lower loneliness in models correcting for age and intelligence (OR = 0.382, 95% CI: 0.167 - 0.777, $p = .012$), and this relationship was robust to multiple testing correction. There were no associations between FP performance and social contact or social support, and these remained non-significant with each step-wise model iteration ($\beta \leq |.176|$, $ps \geq .247$). When including childhood intelligence, higher FP performance remained significantly associated with smaller close social network size in all model iterations ($\beta$ range = $-.574$ to $-.475$, $p$-value range = $< .001$ to $.002$). However, the models predicting loneliness were no longer significant (OR $\leq .742$, $ps \geq .072$) (Supplemental Table 8).

- Insert Table 2 around here -

**Discussion**

This study focused on whether ToM, measured using the FP test, was related to
psychosocial characteristics in 90 healthy older men. It further explored whether early-life
cognitive ability influences the relationship between ToM and psychosocial characteristics in
older age or whether there are additional influences from later-life cognitive, affective,
personality, and socio-demographic factors. In bivariate analyses, better FP performance was
associated with higher scores on measures of \( g_c \), \( g_f \), and the personality traits of openness and
emotional stability, as well as fewer depressive symptoms. When all potential predictors of
FP performance were considered in a simultaneous regression analysis, \( g_c \), \( g_f \), and depressive
symptoms made significant unique contributions – but only \( g_c \) survived correction for
multiple testing. However, neither \( g_c \) nor \( g_f \) predicted FP performance when childhood IQ was
considered, suggesting that lifelong, stable differences in cognitive ability underlie the
relationships between ToM and later-life intelligence.

We also found that individuals more able to detect social slips reported having smaller
close social networks and being less lonely. Importantly, those children with lower
intelligence grew up to have lower FP scores and greater loneliness; however, it appears that
childhood intelligence did not contribute to the relationship between higher FP scores and
smaller close social network size.

The association between FP performance and close social network size suggests that
individuals with greater social-cognitive aptitude build smaller close social circles, or that
poorer social reasoning relates to maintaining a higher number of close others (at an age when
social relationships generally reduce; Lang & Carstensen, 1994). Whereas those individuals
with lower childhood intelligence were at risk of loneliness and poorer FP performance,
childhood intelligence did not account for links between FP scores and network size. Early-
life intelligence does not appear to have a significant impact on the ability to successfully
prune one’s network to maximize emotional satisfaction (Carstensen, 2006).
Another finding was the significant link between FP performance and $g_c$ and $g_f$. In simple bivariate analyses, both were associated with FP performance with comparable magnitudes; the FP-$g_f$ association supports prior findings (Ibanez et al., 2013; Cox et al., 2014), but the FP-$g_c$ relationship had not been examined. Here, when $g_f$ was modelled simultaneously alongside other covariates, $g_c$ remained a significant FP predictor, corroborating prior evidence that verbal ability uniquely moderates various indices of ToM (Meinhardt-Injac et al., 2018). It is also relevant to reports that reading literary fiction may enhance social reasoning in adults (Kidd & Castano, 2013). Our study’s advantage was that it was able to demonstrate that the relationship between intelligence and ToM begins in childhood and persists throughout life.

FP performance was also correlated with self-reported depressive symptom, as measured by the HADS. This is in line with research indicating that current depressive symptoms or diagnosed MDD debilitate ToM regardless of age (Bora & Berk, 2016). Moreover, depressive symptoms were a nominally significant predictor of FP performance beyond the variance this task shares with intelligence itself.

Finally, openness and emotional stability, measured using IPIP, were the Big-Five personality dimensions significantly correlated with FP performance. A similar relationship between openness and ToM has been reported in younger age (Mar, Oatley, & Peterson, 2009; but see Nettle & Liddle, 2008). Higher openness might relate to greater cognitive ability (Soubellet & Salthouse, 2013), which is supported by our finding that the association between openness and FP performance became non-significant when accounting for intelligence. Similarly, emotional stability was no longer significantly linked with FP performance when entered into this regression model, which also included depressive and anxious symptomatology. Future research should further address how personality relates to ToM using various tasks while accounting for intelligence and affect.
Our study has several limitations. Importantly, we used only one ToM measure. Additionally, it would be optimal to determine the influence of psychosocial characteristics on FP performance compared to control stories. However, participants performed almost entirely at ceiling on these, which is common in studies using the FP test (Stone et al., 1998; MacPherson et al., 2002). This indicates that individual FP variability is unlikely to be attributable to participants’ lack of story comprehension. Our social network size data were cross-sectional and so we cannot directly address the phenomenon of network size pruning. Moreover, two measures were based on single items (loneliness and close social network size) and some were binary (loneliness, relationship status, and living arrangement), reducing fidelity to capture individual variation. Also, loneliness was unbalanced and underpowered, limiting the reliability to estimate its predictors. Furthermore, inclusion of solely male participants within a narrow old-age range, and the restricted nature of this cohort study (Johnson, Brett, Calvin, & Deary, 2016), limit the generalizability of our results to females, other ages, and perhaps less healthy participants of the same age. Finally, while our study was sufficiently powered to reliably detect large- and medium-sized effects, our sample size precluded the reliable estimation of small effects and may have resulted in overfitting and inflation of effect sizes (Ioannidis, 2008).

Notwithstanding these limitations, we provide evidence that ToM, or at least the FP test, may be selectively sensitive to psychosocial characteristics in older age; those more aware of social slips were more likely to report fewer close others, and to report being less lonely, but were not more or less socially isolated with respect to social contact or support. Future work might consider whether ToM contributes to the investment in maintaining closer, more positive relationships as individuals age. However, our work is preliminary and should be interpreted with caution until replicated in wider and more varied samples.
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Table 1. Correlation matrix of the study variables

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<td>.001</td>
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<td>.934</td>
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Note. Pearson’s rs (upper diagonal) and ps (lower diagonal) are reported for associations within the Faux Pas subsample (N = 90). MHT = Moray House Test; Close social net. size = close social network size; g_c = crystallized intelligence; g_f = fluid intelligence.
Table 2. Regression results for Faux Pas performance as a predictor of psychosocial characteristics

<table>
<thead>
<tr>
<th>Social contact</th>
<th>β</th>
<th>SE</th>
<th>t</th>
<th>p</th>
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<tbody>
<tr>
<td>FP</td>
<td>−.122</td>
<td>.107</td>
<td>−1.14</td>
<td>.259</td>
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<td>FP + age</td>
<td>−.123</td>
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<td>−1.15</td>
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<tr>
<td>FP + age + intelligence</td>
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<td>−1.01</td>
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<td>FP + age + intelligence + affect</td>
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<td>−1.17</td>
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<td>.156</td>
<td>−1.03</td>
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<td>−0.85</td>
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<td>FP + age</td>
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<td>.116</td>
<td>0.21</td>
<td>.832</td>
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<td>FP + age + intelligence</td>
<td>.047</td>
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<td>0.31</td>
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<td>.151</td>
<td>−0.91</td>
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<tr>
<td>FP + age + intelligence + affect + personality + socio-demography</td>
<td>−.139</td>
<td>.153</td>
<td>−0.91</td>
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<table>
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<th>Close social network size</th>
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<td>FP</td>
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<td>FP + age + intelligence</td>
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<table>
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<td>0.397</td>
<td>1.054</td>
<td>.072</td>
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<td>0.112</td>
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<td>FP + age + intelligence + affect + personality + socio-demography</td>
<td>0.329</td>
<td>0.098</td>
<td>0.884</td>
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</table>

Note. β = standardized coefficients reported except for loneliness (odds ratios); FP = Faux Pas performance; intelligence = gI and gC; affect = symptoms of depression and anxiety; personality = extraversion, agreeableness, conscientiousness, emotional stability, and openness; socio-demography = relationship status and local deprivation. Bold typeface denotes significant p-values following False Discovery Rate correction across all results reported in the table. All variance inflation factors ≤ 2.830. * No longer significant (p = .076) when controlling for childhood intelligence.