Association of apolipoprotein E with intracerebral hemorrhage risk by race/ethnicity

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Apolipoprotein E and Intracerebral Hemorrhage: a Trans-Ethnic Meta-Analysis

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Key Points

**Question:** What is the effect of history of hypertension and Apolipoprotein E (APOE) on intracerebral hemorrhage (ICH) risk in subjects stratified by self-reported race/ethnicity?

**Findings:** In this case control study that included 13,124 adults, having a copy of *APOE* ε4 increased the risk for lobar ICH only in whites, but after propensity score-matching for hypertension burden, Hispanics subjects showed the same effect of *APOE* ε4.

**Meaning:** *APOE* ε4 is confirmed to be a risk factor for lobar ICH. Its effect is present in non-white populations but is masked by differential hypertension burden.
Abstract

Importance
Genetic studies of intracerebral hemorrhage (ICH) have focused mainly on white participants. Genetic risk may vary or could be concealed by differing non-genetic co-exposures in non-white populations. Trans-ethnic analysis of risk may clarify the role of genetics in ICH risk across populations.

Objective
We sought to determine whether established differences in ICH risk by race and ethnicity could be due to variability in the effects of Apolipoprotein E (APOE) epsilon (ε) alleles, the most potent genetic risk factor for ICH.

Design, Setting and Participants
We meta-analyzed the effects of APOE allele status on ICH risk, applying a two-stage clustering approach based on race/ethnicity and contributing study. A propensity score analysis was used to model the influence of APOE against the burden of hypertension across races/ethnicities. Primary ICH cases and controls were collected from hospital- and population-based studies from US and European sites within the International Stroke Genetic Consortium, enrolled from 1999 to 2017. Secondary causes of ICH were excluded from enrollment. Controls were regionally matched within each participating study.
Clinical variables were systematically obtained from structured interview within each site. *APOE* genotype was centrally determined for all studies.

**Results**

13,124 subjects (54.5% male, median age 66 (18-100) years) were included. In whites, *APOE* ε2 (odds ratio (OR)=1.49, 95% confidence interval (CI)=1.24-1.80, p<0.001) and *APOE* ε4 (OR=1.51, 95% CI=1.23-1.85, p<0.001) were associated with lobar ICH risk, however within self-identified Hispanics and blacks, no associations were found. After propensity score-matching for hypertension burden, *APOE* ε4 was associated with lobar ICH risk among Hispanics (OR=1.14, 95% CI=1.03-1.28, p=0.01), but not in blacks (OR=1.02, 95% CI=0.98-1.07, p=0.251). *APOE* ε2 and ε4 did not show an effect on non-lobar ICH risk in any race/ethnicity.

**Conclusions and Relevance**

*APOE* ε4 and ε2 alleles affect lobar ICH risk variably by race and ethnicity. Associations are confirmed in whites but can be shown in Hispanics only when the excess burden of hypertension is propensity score-matched. Further studies are needed to explore interactions between *APOE* alleles and environmental exposures that vary by race and ethnicity in representative populations at risk for ICH.
INTRODUCTION

Spontaneous intracerebral hemorrhage (ICH) is the most severe form of stroke. In the US alone, 160,000 people suffer from ICH each year with a case fatality rate of 54% at 1 year\(^1\). ICH prevalence has increased 47% between 1990 and 2010\(^2\), and ICH risk appears to vary among white, black and Hispanic populations\(^3\)-\(^6\). Compared to whites, young and middle-aged blacks have an almost twofold increased risk for ICH\(^3\),\(^4\). Similarly, Hispanics have a relative risk increase that ranges from 1.4 to 3.7 for lobar and non-lobar ICH, respectively\(^5\). Moreover, not only is hypertension prevalence among the elderly lower among non-Hispanic whites (76.3%) than non-Hispanic blacks (82.5%), or Hispanics (79.2%), but the risk of ICH in the presence of hypertension increases more than 50% from whites to Hispanics\(^7\)-\(^9\). The contributions of genetic and acquired ICH risk factors to these observed risk differences are poorly understood.

Prior studies conducted in predominantly European-ancestry populations have demonstrated that Apolipoprotein E (APOE) \(\varepsilon2\) and \(\varepsilon4\) alleles potently increase risk of lobar ICH\(^10\). In Alzheimer’s disease (AD), another disorder strongly associated with APOE \(\varepsilon\) allele status, the degree of risk contributed by APOE genotype is highly correlated with the ancestry of the population studied. Among non-Hispanic whites, homozygous carriers of APOE \(\varepsilon4\) exhibit up to 12 times higher risk of AD, but this same haplotype exerts little or no effect among blacks or Hispanics\(^11\)-\(^13\).
Understanding how genetic risk factors vary across race and ethnicity may highlight novel underlying disease mechanisms and identify populations who may be particularly responsive to specific prevention strategies, as has previously been shown in treatment response for heart failure by race and ethnicity\textsuperscript{14}. Unfortunately, with individuals of African American and Hispanic ancestry representing less than 4% of all samples in genome-wide association studies (GWAS), only recently has it become possible to study genetic risk of common disease across representative US populations\textsuperscript{15}.

We tested the effects of \textit{APOE} ε alleles on risk of lobar and non-lobar ICH among whites, blacks, and Hispanics, using direct genotyping data supplemented by genome-wide genotyping where available in cases and controls from the International Stroke Genetics Consortium (ISGC). Because these analyses revealed substantial heterogeneity by race and ethnicity, we further explored the degree to which differential burden of hypertension across populations contributes to the variability in observed \textit{APOE} effects.
METHODS

Participating Studies and Data Collection

Case and control subjects included in the study were gathered from 3 multicenter studies in the US and from 8 European sites participating in the ISGC, based on availability of directly ascertained APOE ε genotypes and a harmonized local acute case recruitment scheme. ICH cases from population-based cohorts were not included due to potential imbalances in lethal case recruitment between the two sampling approaches16. Studies included The Genetics of Cerebral Hemorrhage with Anticoagulation (GOCHA) study17, the Genetic and Environmental Risk Factors for Hemorrhagic Stroke (GERFHS) study18, the Ethnic/Racial Variations of Intracerebral Hemorrhage (ERICH) study19, the Hospital del Mar and Vall d’Hebron Hospital ICH studies20,21, the Jagiellonian University Hemorrhagic Stroke Study22, the Lund Stroke Register study23, the Edinburgh Stroke Study and LINCHPIN24, the UMC Utrecht ICH study, and the Brescia Stroke Registry25. Because of variable sample sizes from contributing centers, data from European studies were analyzed together for association testing in meta-analysis (ISGC Europe), as done previously26,27. More specific inclusion and exclusion criteria for each of the included studies are reported in eTable 1. Demographic variables, including self-identified race and ethnicity8 were systematically obtained from structured patient and family member interview within each site19,28, along with additional covariates29. CT images on admission were analyzed at each participating site for classification as lobar (involving predominantly
the cortex and underlying white matter), and non-lobar (involving predominately the basal
ganglia, periventricular white matter, or internal capsule), following prespecified criteria
26,27. \textit{APOE} genotype was centrally determined following standard procedures\textsuperscript{30}. Genome-
wide data were available for a subgroup of subjects. Genetic and bioinformatic analysis
followed standardized prespecified quality control procedures\textsuperscript{31} (see supplementary
methods). IRB approval was obtained at all participating centers, and informed consent
was obtained from all participants or their legally authorized representative.

**Population Stratification**

Fifteen ancestry informative markers were selected from subjects with direct or genome-
wide genotyping and subjected to principal component analysis in accordance with
previously published methods\textsuperscript{32–35}. The first four principal components were included in
regression models to adjust for population stratification in this subgroup. This PC analysis
was not used to reclassify participants, as self-identified race/ethnicity may capture
exposures that transcend genetic ancestry and could contribute to explain the stratification
among different populations. A complete description of the genetic analysis, the subjects
genotyped, and the markers selected is available in the supplements (eTable 2).

**Statistical Analysis**

Categorical variables were expressed as count (%), and continuous variables as median
(interquartile range, IQR) or mean (standard deviation, SD), as appropriate. Categorical
variables were compared using the χ² test whereas continuous variables were compared with Mann-Whitney tests.

We tested APOE allele association with ICH risk using three logistic regression models. Model 1 was adjusted for age, sex, and history of hypertension. Model 2 included variables from Model 1 in addition to history of hypercholesterolemia, history of previous ischemic stroke, warfarin, statin and antiplatelet use, smoking and alcohol use. Model 3 also included variables from Model 1 and added the first four principal components (PCs) derived from ancestry-informative genotypes. APOE risk allele status was modeled as two variables, ε2 and ε4 coded for allele counts (0, 1, or 2 for each) in an additive model referent to the wildtype ε3 allele. Analyses were performed in lobar and non-lobar ICH, given the known differences in underlying biology between the two ICH locations. All statistical analyses were performed using STATA (Version 13.0; Stata Corp) and R statistical software (http://www.r-project.org).

Trans-ethnic Meta-analysis

We applied a two-stage clustering approach for meta-analysis, based on race/ethnicity and stratified by study. Cases and controls in each study were divided into blacks, whites, and Hispanics, based on self-identified race/ethnicity. Each race/ethnicity group within each study were allocated to the same cluster and tested using regression models described above. Individual cluster results were presented graphically by plotting odds ratio estimates on a Forest plot to visually assess heterogeneity. The effects sizes obtained were then used
for a Der Simonian-Laird random effects, inverse-weighted non-parametric meta-analysis\textsuperscript{38}. Cochrane’s Q and $I^2$ were used to quantify heterogeneity.

Propensity Score Modeling of $APOE$ and Hypertension

To address imbalances in the burden of hypertension across ICH populations, and related imbalances of baseline characteristics among subjects with and without hypertension, two propensity score (PS) analyses were performed using the nearest neighbor matching method to compare subjects of similar underlying hypertensive pathophysiology burden\textsuperscript{39,40}. The first PS analysis was constructed based on history of hypertension, and included variables of age, sex, and self-identified ethnicity/race. The second PS analysis, leveraging data only available in the ERICH study, contained the same variables as the first PS analysis, in addition to the number of medications prescribed to treat hypertension, and systolic and diastolic blood pressure at ICH presentation. Propensity score results were used in a logistic regression model for ICH risk identical to Model 1 described above. As a sensitivity analysis, the same propensity score procedure was tested against age (greater than or less than 65 years), sex, and hypercholesterolemia, to increase the confidence that the PS findings were specific for hypertension.

Power Calculation

Using empiric data from our analyses, we performed a post-hoc calculation of our statistical power to detect an effect of $APOE \varepsilon$ alleles on lobar ICH risk in blacks and Hispanics commensurate with the effect size detected in whites. Type I error rate of 0.05,
log additive inheritance mode, and 0.01 of population risk were assumed, with analyses performed using Quanto software version 1.2.4 (http://hydra.usc.edu/gxe)⁴¹.
RESULTS

13,124 subjects (47.2% cases) were included from the participating studies. 54.5% were males, median age was 66 years (inter-quartile range (IQR): 56-76), with 8,334 whites, 2,273 blacks, 1,781 Hispanics, and 736 subjects of other race/ethnicity (Table 1). The latter were excluded from the primary analyses given low statistical power. Rates of \( APOE \varepsilon4 \) homozygosity in cases were 3.6%, 5.3%, and 1.8%, whereas rates of \( APOE \varepsilon2 \) homozygosity in cases were 1.0%, 1.2%, and 0.4% respectively in whites, blacks and Hispanics. 56.8% subjects (4,069 cases and 3,379 controls) had genome-wide or direct genotyping data on ancestry informative markers for PC analysis (eTable 3). Self-identified race and ethnicity showed overall strong concordance with PC-based ancestry (eFigure 1). Additional clinical covariates were available for a subset of subjects (eTable 4).

Lobar ICH

We analyzed 2,305 lobar ICH cases from all studies. Model 1 confirmed the previously reported effect of \( APOE \varepsilon2 \) (pooled odd ratio (OR) = 1.49, 95% confidence intervals (CI) = 1.24-1.80, \( p < 0.001 \)) and \( APOE \varepsilon4 \) (pooled OR = 1.51, 95% CI = 1.23-1.85, \( p < 0.001 \)) on ICH risk, however within self-identified Hispanics and blacks, no associations were found (Figure 1). Model 2 was used to interrogate the independent effect of \( APOE \) alleles on ICH, controlled for established ICH predictors (Figure 2). Here, \( APOE \varepsilon2 \) and \( \varepsilon4 \) allele retained association with lobar ICH. As with Model 1, this effect was observed in whites, but not in blacks or Hispanics (for \( APOE \varepsilon2 \), OR =1.45, 95% CI = 1.04-2.03, \( p = 0.028; \)
for \textit{APOE} ε4, OR = 1.51, 95% CI = 1.14-1.99, p = 0.004). Model 3 considered population stratification (Figure 3). In whites, both \textit{APOE} ε2 (OR = 1.81, 95% CI = 1.33-2.45, p < 0.001) and \textit{APOE} ε4 (OR = 1.80, 95% CI = 1.33-2.44, p < 0.001) conferred higher risk for lobar ICH. For \textit{APOE} ε4 alone we found a similar effect in Hispanics, suggesting that population stratification may have played some role in the lack of ε4 effect found in Models 1 and 2, particularly for the large and ethnically diverse Hispanic population recruited through the ERICH study. In contrast, for blacks neither \textit{APOE} ε2 nor \textit{APOE} ε4 conferred a significant risk for lobar ICH after controlling for population structure.

**Non-lobar ICH**

We analyzed 3,897 non-lobar ICH cases (Figure 1). In Model 1, \textit{APOE} ε2 and ε4 did not show an effect on non-lobar ICH risk, across any of the self-identified race/ethnicity groups. When comparing non-lobar ICH cases vs controls, \textit{APOE} ε4 p-values were 0.219 for whites, 0.182 for blacks, and 0.346 for Hispanics. For Model 2 and model 3 in non-lobar ICH, again neither \textit{APOE} ε2 nor \textit{APOE} ε4 showed an association with disease risk across all the studies and races/ethnicities (Figures 2 and 3).

**Power calculation (lobar ICH):**

Given the differences in sample sizes between whites, blacks, and Hispanics, we performed post-hoc power calculations to determine whether our study was powered to detect a comparable \textit{APOE} effect in the smaller populations of blacks and Hispanics. Given the frequency of \textit{APOE} ε4 in blacks (ε4 frequency 37.7%), our sample size (assuming an
unmatched case-control ratio of 1:1) would provide 99% power to detect an ε4 effect similar to the lower bound of the 95% CI seen in whites (OR=1.43). Our analyses of APOE ε4 effects in Hispanics were similarly powered at 90%. Further, the APOE ε2 frequency in blacks (19.9%) at the reported sample sizes would provide 93.8% power to detect the lower bound of the effect seen in whites (OR=1.38). For Hispanics, given the lower APOE ε2 frequency (0.8%) 80% power would be achieved at a slightly higher effect size (OR=1.60), but still below the one found in whites.

Propensity Score Modeling for Hypertension

We used a propensity score (PS) analysis to attempt to isolate the influence of APOE against the imbalanced burden of hypertension across race/ethnicity. In our first PS, we selected case and control subjects with a balanced hypertension burden, comprised of individuals of white, black, and Hispanic ancestry. In this matched and homogeneous group, we were able to detect an effect of APOE ε4 on lobar ICH risk among Hispanics (OR = 1.14, 95% CI = 1.03-1.28, p = 0.01), but not in blacks (OR = 1.02 95% CI = 0.98-1.07, p = 0.251). Results were confirmed in our secondary PS analysis performed only in the ERICH dataset, which included hypertension diagnosis as well as additional hypertension severity variables including number of medications used to treat hypertension, and systolic and diastolic blood pressure readings (Figure 4).
DISCUSSION

Although \textit{APOE} effects on ICH risk have been characterized in multiple prior studies and meta-analyses for European, and more recently Asian, ancestries, there have been fewer opportunities for examination of US minority populations at disproportionate risk for ICH. Supplemented by data from the ERICH study, we are now able to confirm variability in associations between \textit{APOE} $\epsilon$ genotypes and lobar ICH risk across whites, blacks, and Hispanics, and explore the degree to which differences in genetic risk are attributable to comorbid exposures. Our results demonstrate an effect of \textit{APOE} $\epsilon 4$ and $\epsilon 2$ alleles in lobar ICH led primarily by white individuals and confirmed by additional models adjusting for known covariates	extsuperscript{29}. When the effect of hypertension is propensity-matched across race and ethnicity, \textit{APOE} $\epsilon 4$ emerges as a risk factor for lobar ICH among self-identified Hispanic individuals.

Our results highlight the challenges in generalizing genetic risk factors across ancestries, where non-genetic exposures are known to vary by race and ethnicity. In AD, the relative risks for Hispanics or blacks associated with an \textit{APOE} $\epsilon 4$ allele become progressively weaker or disappear entirely in comparison to whites	extsuperscript{42–46}. In ICH, \textit{APOE} $\epsilon$ alleles have already been shown to exert higher effects in East Asians when compared to subjects of European ancestry	extsuperscript{47}. While recent analyses by Sawyer et al.	extsuperscript{48} demonstrate the effect of hypertension and \textit{APOE} $\epsilon$ allele status on ICH risk across race/ethnicity specific to the ERICH study, the present analysis benefits from a larger sample size via formal trans-
ethnic meta-analysis as well as a propensity score matching approach that helps to illustrate the potential mechanisms underlying the observed variability of $APOE \varepsilon$ alleles on lobar ICH risk across populations.

It is important to note that the observed differences in association between $APOE \varepsilon$ alleles and lobar ICH risk do not provide direct evidence that biological effects of the $APOE$ gene or associations with underlying cerebral amyloid angiopathy (CAA, a major cause of lobar ICH) are necessarily different across racial and ethnic boundaries. It would seem more likely that genetic and/or environmental risk exposures covarying with race and ethnicity exert a role in modifying or mitigating underlying $APOE$ genetic effects. Our propensity score analysis supports this conjecture, demonstrating that hypertension, the most important known risk factor for ICH, may simply obscure $APOE$ effects that may indeed be common across ancestries. Aside from variation in environmental risk exposures, variants in a modifier gene (or genes) that differ across populations may alter the biological effect of $APOE$ and consequently vary ICH risk, as has been hypothesized for AD$^{49,50}$. Furthermore, genetic variants that are racially stratified and not related to $APOE$ may directly modify the risk of ICH. This hypothesis represents an alternative explanation for why the propensity score matching for hypertension only partially remediated the effect of $APOE \varepsilon4$ on ICH risk in Hispanics, and had little to no effect in blacks. $APOE$ interaction studies and trans-ethnic GWAS for ICH will likely provide insights on these hypotheses. Similarly, analyses of CAA in non-white populations will additionally clarify the effect of race/ethnicity on this pathological pathway.
In this study, we have not attempted to stratify subsets of subjects by probable CAA status using MRI data as has been done in prior meta-analyses. In fact, most prior studies linking lobar hemorrhage locations to the pathological diagnosis of CAA were performed in largely white populations\textsuperscript{52}. As such, widely accepted criteria for classifying probable and possible CAA using hemorrhage location and microbleed counts have not been validated in non-white populations\textsuperscript{53}. Validating CAA burdens across multi-ethnic populations will require concomitant neuroimaging and/or tissue pathology data in genotyped individuals of many races and ethnicities to ensure patients are not mis-assigned.

Previously demonstrated associations between $APOE$ $\varepsilon 4$ and non-lobar ICH risk, also seen in non-lobar ICH recurrence\textsuperscript{17,54}, were not replicated in this study. Potential explanations include a higher rate of subjects affected by hypertension and an overall younger age of subjects in our study. These factors may reflect the driving effects of environmental or non-$APOE$ genetic exposures in younger populations with non-lobar ICH in particular. Demographic heterogeneity is also higher in our study and the reduced availability of covariates such as steady-state lipid levels\textsuperscript{54} for risk modeling may have influenced this finding. Finally, our previous meta-analysis of $APOE$ effects in ICH also failed to show the association between $\varepsilon 4$ and non-lobar ICH in blacks, a finding supported by the present analyses\textsuperscript{17}. Future studies in larger datasets with well-phenotyped cases will be needed to further elucidate the potential role of $APOE$ $\varepsilon 4$ in non-lobar ICH.
The targeted enrollment of Hispanics and blacks through the ERICH study, lacking in prior reports\textsuperscript{17,55}, is an important strength of the present study. The high number of non-whites enrolled permits well-powered analyses in these populations and promotes confidence that the lack of observed effects is not a false acceptance of the null hypothesis, as supported by our post hoc power calculations. Some limitations should be acknowledged. Diagnosis of comorbidities are based on self-identified attestation and are therefore influenced by patient or caregiver awareness. This concern is present in both cases and controls, however, and internal consistency between diagnoses and prescribed medications helps to limit this potential source of bias. Furthermore, our propensity score is based on variables only partially capturing the complex phenotype represented by hypertension. However, this lack of information content is likely to bias our score results towards the null; we expect that a more precise index of hypertension burden would have increased our ability to normalize this phenotype across ethnicities and demonstrate an even more homogeneous \textit{APOE} \(\varepsilon4\) effects. Finally, genome-wide genotypes for ERICH participants are not currently available, preventing us from determining whether additional genetic exposures modify the role of \textit{APOE} on ICH risk across race/ethnicity.

In this meta-analysis, \textit{APOE} \(\varepsilon2\) and \(\varepsilon4\) remain genetic risk factors for lobar ICH but these results are largely driven by the strong associations in white individuals. However, our results support a biological effect of \textit{APOE} \(\varepsilon4\) alleles that seems to transcend ancestral backgrounds\textsuperscript{56}, albeit with varying effect due to the presence of racial and ethnic disparities.
across related risk factors. As availability of genetic data on US minority populations continues to increase, it is hoped that improved modeling of covarying genetic and non-genetic exposures in these populations will provide new insights into treatment and prevention strategies in ICH that maximize the potential benefits for all.
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1) SM had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

2) conception and design of the study (SM; CDL; DW; JR; CDA)

3) acquisition and analysis of data (KC; AM; MJL; AP; CJM; MLF; JM; JR; JJ; EG; RE; ECG; CS; AS; JMJ; JP; AU; AP; AL; BMH; JLM; DLT; MS; DLB; SLS; BBW; JFM; CSK; FDT; SJK; HS; CE; IJD; KR; NS; RAS; CLS; CJKMK; KMN; IFC; PD; BN; JNG; AV; SMG; GJF; AB; CDL; DW; JR; CDA)

4) drafting a significant portion of the manuscript or figures (SM; KR; AV; AB; GJF; CJKMK; CDA; JR)

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POTENTIAL CONFLICTS OF INTEREST:

Dr. Anderson is supported by grants from the NIH, the American Heart Association, and the Massachusetts General Hospital Center for Genomic Medicine, and has consulted for ApoPharma, Inc.

Prof Klijn is supported by a clinical established investigator grant from the Dutch Heart Foundation (grant number 2012T077), and an Aspasia grant from Zonmw (grant number 045008048).
Table 1: Demographic characteristics, clinical and APOE allele frequencies across participating studies.

<table>
<thead>
<tr>
<th>STUDIES</th>
<th>ERICH n=5017</th>
<th>GOCHA n=2297</th>
<th>ISGC Europe n=3471</th>
<th>GERFHS n=2339</th>
</tr>
</thead>
<tbody>
<tr>
<td>Male sex, n (%)</td>
<td>2866 (57.1)</td>
<td>1266 (55.1)</td>
<td>1891 (54.6)</td>
<td>1130 (48.3)</td>
</tr>
<tr>
<td>Age, median (IQR)</td>
<td>61 (52-72)</td>
<td>73 (65-80)</td>
<td>70 (61-77)</td>
<td>65 (51-75)</td>
</tr>
<tr>
<td>Cases, n (%)</td>
<td>2880 (57.4)</td>
<td>1322 (57.6)</td>
<td>1281 (36.9)</td>
<td>811 (34.7)</td>
</tr>
<tr>
<td>Lobar ICH, n (%)</td>
<td>882 (30.6)</td>
<td>613 (47.8)</td>
<td>493 (40.2)</td>
<td>316 (39.0)</td>
</tr>
<tr>
<td>Non-lobar ICH, n (%)</td>
<td>1998 (69.4)</td>
<td>670 (52.2)</td>
<td>734 (59.8)</td>
<td>495 (61.0)</td>
</tr>
<tr>
<td>Hypertension, n (%)</td>
<td>3364/4976 (67.6)</td>
<td>1667/2275 (73.3)</td>
<td>1673/2893 (57.8)</td>
<td>1264/2337 (54.1)</td>
</tr>
<tr>
<td>Self-reported race/ethnicity, n (%)</td>
<td>white 1739 (34.7)</td>
<td>2024 (88.1)</td>
<td>2622 (75.5)</td>
<td>1949 (83.3)</td>
</tr>
<tr>
<td></td>
<td>blacks 1751 (34.9)</td>
<td>131 (5.7)</td>
<td>-</td>
<td>390 (16.7)</td>
</tr>
<tr>
<td></td>
<td>Hispanics 1527 (30.4)</td>
<td>60 (2.6)</td>
<td>194 (5.6)</td>
<td>-</td>
</tr>
<tr>
<td></td>
<td>other/missing -</td>
<td>82 (3.6)</td>
<td>654 (18.8)</td>
<td>-</td>
</tr>
<tr>
<td>APOE ε4 allele count, n (%)</td>
<td>0 3553 (70.8)</td>
<td>1664 (71.8)</td>
<td>2789 (80.4)</td>
<td>1664 (71.1)</td>
</tr>
<tr>
<td></td>
<td>1 1298 (25.9)</td>
<td>570 (24.8)</td>
<td>637 (18.4)</td>
<td>601 (25.7)</td>
</tr>
<tr>
<td>APOE ε2 allele count, n (%)</td>
<td>0</td>
<td>1</td>
<td>2</td>
<td></td>
</tr>
<tr>
<td>----------------------------</td>
<td>---</td>
<td>---</td>
<td>---</td>
<td></td>
</tr>
<tr>
<td></td>
<td>4262 (85.0)</td>
<td>1916 (83.4)</td>
<td>3034 (87.4)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>710 (14.2)</td>
<td>363 (15.8)</td>
<td>413 (11.9)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>45 (0.9)</td>
<td>18 (0.8)</td>
<td>24 (0.7)</td>
<td></td>
</tr>
</tbody>
</table>
FIGURE LEGENDS:

Figure 1:
Forest plots of meta-analysis of APOE in lobar and non-lobar ICH cases and controls in Model 1, stratified across participating studies and race/ethnicity.

Figure 2:
Forest plots of meta-analysis of APOE in lobar ICH and non-lobar ICH cases and controls, in Model 2 (adjusted for age, sex, history of hypertension, hypercholesterolemia, warfarin, statin and antiplatelet use, smoking and alcohol use), stratified across participating studies and race/ethnicity.

Figure 3:
Forest plots of meta-analysis of APOE in Lobar and Non-lobar ICH cases and controls, in Model 3 (adjusting for Model 1 covariates plus principal components 1 and 2), stratified across participating studies and race/ethnicity.

Figure 4:
Risk of APOE ε4 allele for lobar ICH, across different race/ethnicities after propensity score matching based on hypertension burden.
REFERENCES:


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doi:10.1001/archneur.63.3.431.

doi:10.1016/j.jalz.2014.06.015.


22. Pera J, Slowik A, Dziedzic T, Pulyk R, Wloch D, Szczudlik A. Glutathione peroxidase 1


Supplementary Online Content


eMethods

eTable 1. ICH Case Inclusion and Exclusion Criteria by Recruitment Site

eTable 2. Ancestry Informative Markers Selected

eTable 3. Comparison Between With Subjects With and Without Ancestry Informative Markers Available for Population Stratification Analysis

eTable 4. Frequencies of Additional Diseases and Risk Factors Among Cases and Controls of the Available Data

eFigure. Distribution of Subjects in Principal Component (PA) Space and the Relation with Self-identified Race and Ethnicity (Ellipses Cluster 95% of PC Analysis Data)

This supplementary material has been provided by the authors to give readers additional information about their work.
Apolipoprotein E and Intracerebral Hemorrhage: A Trans-Ethnic Meta-Analysis

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eMethods
Participating Studies
Case and control subjects included in the study were gathered from 3 multicenter studies in the US and from 8 distinct European sites participating in the ISGC, based on availability of directly ascertained APOE ε genotypes. US studies included The Genetics of Cerebral Hemorrhage with Anticoagulation (GOCHA) study, the Genetic and Environmental Risk Factors for Hemorrhagic Stroke (GERFHS) study, the Ethnic/Racial Variations of Intracerebral Hemorrhage (ERICH) study. European studies included The Hospital del Mar and Vall d’Hebron Hospital ICH studies in Barcelona, Spain, the Jagiellonian University Hemorrhagic Stroke Study in Krakow, Poland, the Lund Stroke Register study in Lund, Sweden, the Edinburgh Stroke Study and LINCHPIN study in Scotland, UK, the University Medical Center (UMC) Utrecht ICH study, and the Brescia Stroke Registry. Because of variable sample sizes from contributing centers, data from the European studies (ISGC Europe) were analyzed together for association testing in meta-analysis, as done previously. We analyzed primary ICH cases, CT or MRI confirmed, aged >18 years (GOCHA enrolled subjects aged > 50 years). ICH cases where there was evidence of secondary cause (such as trauma, tumor, hemorrhagic transformation of ischemic stroke, or vascular malformation) were excluded. More specific inclusion and exclusion criteria for each of the included studies are reported in eTable 1.

Phenotypic and Genetic variables
Demographic variables, including self-identified race and ethnicity (according to the recommendation of the Standards for Maintaining, Collecting, and Presenting Federal Data on Race and Ethnicity\textsuperscript{12}) were systematically obtained from structured patient and family member interview within each site\textsuperscript{3,13}. Additional covariates previously associated with ICH\textsuperscript{14} risk were also collected: medical history of hypertension, previous stroke, hypercholesterolemia, warfarin, antiplatelet and statin use, smoking history, and alcohol exposure (more than twice per week)\textsuperscript{15}. Given the heterogeneity among the studies, other variables were not available. Number of medications used to treat hypertension (taken in the 2 weeks before ICH onset or before the interview, for cases and controls respectively) and systolic and diastolic blood pressure readings (measured on enrollment with index ICH for cases, and at the time of interview with controls) were also available for ERICH participants. CT images on admission were analyzed at each participating site for hematoma classification as lobar (involving predominantly the cortex and underlying white matter), and non-lobar (involving predominately the basal ganglia, periventricular white matter, or internal capsule), following prespecified criteria\textsuperscript{11}.

APOE genotype was determined as part of an ongoing GWAS of ICH. Taqman (Applied Biosystems, Foster City, CA) and iPLEX (Sequenom, San Diego, CA) methodologies were adopted at the University of Miami and Massachusetts General Hospital genotyping centers, respectively for all the sample. The genotypes obtained for rs429358 (C/T) and rs7412 (C/T) were used to define the three standard human APOE $\varepsilon$ haplotypes of $\varepsilon2$, $\varepsilon3$, and $\varepsilon4$\textsuperscript{16}. Genome-wide data were available for a subgroup of subjects, from Affymetrix 6.0 in GERFHS, and Illumina HumanHap610-Quad in GOCHA and ISGC Europe. IRB approval was obtained at all participating centers, and informed consent was obtained from all participants or their legally authorized representative.

**Population Stratification**

Fifteen ancestry informative markers were selected among the polymorphisms directly genotyped in ERICH\textsuperscript{17} (eTable 2). These markers, distributed across the genome, showed difference in allele frequency ($\delta$) of at least 0.5 between any two of the recruited ancestral populations (African American, Hispanic American, non-Hispanic European). GWAS data for GOCHA, GERFHS, ISGC Europe samples were analyzed as previously described\textsuperscript{10}. Briefly
Standardized prespecified quality control procedures\textsuperscript{18}, imputation via IMPUTE2 v.2.2\textsuperscript{19} and 1000 Genomes integrated reference panels (Phase I interim release in NCBI build 37), and post imputation filtering (MAF <0.01, IMPUTE2 information score <0.7) were implemented. Samples with < 90% call rates, or SNPs with call rates <95% or deviation from Hardy-Weinberg equilibrium at P-value < $2.5 \times 10^{-4}$ were excluded. Based on these 15 selected markers, principal component analysis was implemented in GOCHA, GERFHS, ISGC Europe and ERICH subjects in accordance with previously published methods\textsuperscript{20–22}. The first four principal components were included in regression models to adjust for population stratification. We have previously shown the reliability of this approach in defining ancestries of populations\textsuperscript{17}. This PC analysis was not used to reclassify participants, as self-identified race/ethnicity may capture exposures that transcend genetic ancestry and could contribute to explain the stratification among different populations.
### eTable 1. ICH Case Inclusion and Exclusion Criteria by Recruitment Site

<table>
<thead>
<tr>
<th>Study</th>
<th>Inclusion Criteria</th>
<th>Exclusion Criteria</th>
<th>Control Recruitment</th>
</tr>
</thead>
</table>
| Brescia Stroke Registry. (University of Brescia, Brescia, Italy) | (Hospital-based, prospective, +18 y/o)  
  • Acute hospitalization for ICH  
  • CT or MRI confirmation of ICH  
  • Age > 18 | • Head trauma  
  • Brain tumor  
  • Ischemic stroke  
  • Vascular malformation  
  • Other cause of secondary ICH | • Regionally matched, hospital and ambulatory clinics |
| UMC Utrecht ICH Study (University Medical Center Utrecht, Utrecht, The Netherlands) | (Hospital-based, prospective, +18 y/o)  
  • Acute hospitalization for ICH  
  • CT confirmation of ICH  
  • Age > 18 | • Head trauma  
  • Brain tumor  
  • Ischemic stroke  
  • Vascular malformation  
  • Other cause of secondary ICH present on admission or in follow-up | • Regionally matched, blood donor population |
| Edinburgh – ESS (Western General Hospital, Edinburgh, Scotland, UK) | (Inpatient and outpatient hospital-based, prospective, +55 y/o)  
  • Acute hospitalization for ICH  
  • CT or MRI confirmation of ICH  
  • Age > 55 | • Head trauma  
  • Brain tumor  
  • Ischemic stroke  
  • Vascular malformation  
  • Presentation > 1 week from ICH  
  • Antecedent drug use  
  • Primary coagulopathy | • N/A |
| Edinburgh – LINCHPIN (Western General Hospital, Royal Infirmary of Edinburgh, St. John’s Hospital at Howden, West Lothian, Scotland, UK) | (Community-based in areas served by NHS Lothian Health Board, prospective with hot-pursuit and retrospective augmentation, +16 y/o)  
  • Symptomatic ICH (acute or chronic) | • Head trauma  
  • Brain tumor  
  • Ischemic stroke with hemorrhagic transformation  
  • Vascular malformation  
  • Other cause of secondary ICH | • N/A |
<table>
<thead>
<tr>
<th>Study</th>
<th>Inclusion Criteria</th>
<th>Exclusion Criteria</th>
<th>Matched Method</th>
</tr>
</thead>
<tbody>
<tr>
<td>ERICH (19 centers in USA, based at University of Cincinnati)</td>
<td>CT or MRI confirmation of acute or chronic ICH, Age &gt; 16, Resident in area served by NHS Lothian Health Board at time of ICH</td>
<td>Head trauma, Brain tumor, Ischemic stroke, Vascular malformation, Other cause of secondary ICH</td>
<td>Regionally matched, random-digit-dialing</td>
</tr>
<tr>
<td>GOCHA (6 centers in USA, based at Massachusetts General Hospital)</td>
<td>CT or MRI confirmation of acute or chronic ICH, Age &gt; 18</td>
<td>Head trauma, Brain tumor, Ischemic stroke, Vascular malformation, Other cause of secondary ICH</td>
<td>Regionally matched, ambulatory clinics</td>
</tr>
<tr>
<td>GERFHS (16 centers in the Greater Cincinnati/Northern Kentucky region of USA, based at University of Cincinnati)</td>
<td>CT or MRI confirmation of acute or chronic ICH, Age &gt; 18</td>
<td>Head trauma, Brain tumor, Ischemic stroke, Vascular malformation, Other cause of secondary ICH</td>
<td>Regionally matched, random-digit-dialing</td>
</tr>
<tr>
<td>ISGC Europe ICH studies (Hospital del Mar, Vall d’Hebron Hospital, Jagiellonian University, Lund University)</td>
<td>CT or MRI confirmation of acute or chronic ICH, Age &gt; 18</td>
<td>Head trauma, Brain tumor, Ischemic stroke, Vascular malformation, Other cause of secondary ICH</td>
<td>Regionally matched, hospital and ambulatory clinics</td>
</tr>
<tr>
<td>* Lothian Birth Cohort 1936</td>
<td>Community population born in 1936 who took Scottish Mental Health Services Questionnaire</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>
(All centers serving the Lothian Area of Scotland) |   | Survey in 1947, living in Lothian, Scotland, UK

* This study was used to recruit controls for the Edinburgh – ESS and Edinburgh – LINCHPIN cases
**eTable 2. Ancestry Informative Markers Selected**

<table>
<thead>
<tr>
<th>AIM</th>
<th>Chromosome</th>
<th>position (GRCh37.p13)</th>
<th>A1</th>
<th>A2</th>
<th>AFR</th>
<th>AMR</th>
<th>EUR</th>
</tr>
</thead>
<tbody>
<tr>
<td>rs7581299</td>
<td>2</td>
<td>179900720</td>
<td>T</td>
<td>C</td>
<td>0.749</td>
<td>0.660</td>
<td>0.370</td>
</tr>
<tr>
<td>rs11725412</td>
<td>4</td>
<td>38277754</td>
<td>A</td>
<td>G</td>
<td>0.215</td>
<td>0.464</td>
<td>0.063</td>
</tr>
<tr>
<td>rs12640848</td>
<td>4</td>
<td>71506412</td>
<td>A</td>
<td>G</td>
<td>0.967</td>
<td>0.660</td>
<td>0.331</td>
</tr>
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<td>rs1423099</td>
<td>5</td>
<td>75427518</td>
<td>T</td>
<td>C</td>
<td>0.930</td>
<td>0.568</td>
<td>0.264</td>
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<tr>
<td>rs4463276</td>
<td>6</td>
<td>145055331</td>
<td>A</td>
<td>G</td>
<td>0.088</td>
<td>0.376</td>
<td>0.771</td>
</tr>
<tr>
<td>rs13259288</td>
<td>8</td>
<td>4483059</td>
<td>T</td>
<td>G</td>
<td>0.822</td>
<td>0.478</td>
<td>0.764</td>
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<tr>
<td>rs12679427</td>
<td>8</td>
<td>72242674</td>
<td>T</td>
<td>C</td>
<td>0.437</td>
<td>0.424</td>
<td>0.724</td>
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<tr>
<td>rs10962599</td>
<td>9</td>
<td>16795286</td>
<td>T</td>
<td>C</td>
<td>0.977</td>
<td>0.651</td>
<td>0.274</td>
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<tr>
<td>rs10840311</td>
<td>11</td>
<td>9854857</td>
<td>T</td>
<td>C</td>
<td>0.726</td>
<td>0.601</td>
<td>0.352</td>
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<tr>
<td>rs3019657</td>
<td>11</td>
<td>134511647</td>
<td>A</td>
<td>G</td>
<td>0.696</td>
<td>0.496</td>
<td>0.871</td>
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<tr>
<td>rs550338</td>
<td>12</td>
<td>24512037</td>
<td>A</td>
<td>G</td>
<td>0.697</td>
<td>0.460</td>
<td>0.764</td>
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<tr>
<td>rs200354</td>
<td>14</td>
<td>99375321</td>
<td>T</td>
<td>G</td>
<td>0.336</td>
<td>0.476</td>
<td>0.829</td>
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<tr>
<td>rs12913832</td>
<td>15</td>
<td>28365618</td>
<td>A</td>
<td>G</td>
<td>0.972</td>
<td>0.798</td>
<td>0.364</td>
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<tr>
<td>rs2216594</td>
<td>19</td>
<td>33533258</td>
<td>A</td>
<td>G</td>
<td>0.048</td>
<td>0.347</td>
<td>0.676</td>
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<tr>
<td>rs801712</td>
<td>22</td>
<td>47090243</td>
<td>C</td>
<td>G</td>
<td>0.251</td>
<td>0.477</td>
<td>0.795</td>
</tr>
</tbody>
</table>

AIM: ancestry informative marker; A1 and A2: allele frequency; AFR: Afro-American; AMR: Mixed American; EUR: European
**eTable 3. Comparison Between With Subjects With and Without Ancestry Informative Markers Available for Population Stratification Analysis**

<table>
<thead>
<tr>
<th>GENETIC DATA for population stratification</th>
<th>available (n 7452)</th>
<th>not available (n 5672)</th>
<th>p value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Male sex, n (%)</td>
<td>4140 (55.6%)</td>
<td>3013 (53.1%)</td>
<td>0.003</td>
</tr>
<tr>
<td>Age, median (IQR)</td>
<td>66 (54-76)</td>
<td>68 (58-77)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Cases, n (%)</td>
<td>4069 (54.6)</td>
<td>2225 (39.2)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>ICH location</td>
<td></td>
<td></td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Lobar ICH, n (%)</td>
<td>1408/6202 (34.7)</td>
<td>897/6202 (41.9)</td>
<td></td>
</tr>
<tr>
<td>Non-lobar ICH, n (%)</td>
<td>2654/6202 (65.3)</td>
<td>1243/6202 (58.1)</td>
<td></td>
</tr>
<tr>
<td>Hypertension, n (%)</td>
<td>4609/12481 (67.2)</td>
<td>3359/12481 (59.7)</td>
<td></td>
</tr>
<tr>
<td>Self-reported race/ethnicity, n (%)</td>
<td></td>
<td></td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>white</td>
<td>3594/12449 (52.8)</td>
<td>4971/12449 (88.2)</td>
<td></td>
</tr>
<tr>
<td>blacks</td>
<td>1730/12449 (25.4)</td>
<td>550/12449 (9.8)</td>
<td></td>
</tr>
<tr>
<td>Hispanics</td>
<td>1486/12449 (21.8)</td>
<td>118/12449 (2.1)</td>
<td></td>
</tr>
<tr>
<td>APOE ε4 allele count, n (%)</td>
<td></td>
<td></td>
<td>&lt;0.001</td>
</tr>
<tr>
<td></td>
<td>5310 (71.3)</td>
<td>4346 (76.6)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>1909 (25.6)</td>
<td>1197 (21.1)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>233 (3.1)</td>
<td>129 (2.3)</td>
<td></td>
</tr>
<tr>
<td>APOE ε2 allele count, n (%)</td>
<td></td>
<td></td>
<td>0.53</td>
</tr>
<tr>
<td></td>
<td>6251 (83.9)</td>
<td>4842 (85.4)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>1137 (15.3)</td>
<td>780 (13.8)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>64 (0.9)</td>
<td>50 (0.9)</td>
<td></td>
</tr>
</tbody>
</table>
**eTable 4. Frequencies of Additional Diseases and Risk Factors Among Cases and Controls of the Available Data**

<table>
<thead>
<tr>
<th></th>
<th>Controls</th>
<th>Cases</th>
<th>p value*</th>
</tr>
</thead>
<tbody>
<tr>
<td>Smoking, n (%)</td>
<td>2084 (36.2)</td>
<td>1959 (35.3)</td>
<td>0.170</td>
</tr>
<tr>
<td>Alcohol use, n (%)</td>
<td>2414 (47.0)</td>
<td>2087 (41.4)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>Previous CV ischemic event, n (%)</td>
<td>371 (7.7)</td>
<td>636 (12.6)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>Hypercholesterolemia, n (%)</td>
<td>2748 (46.3)</td>
<td>2315 (40.1)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>Statin use, n (%)</td>
<td>1209 (26.6)</td>
<td>1220 (22.7)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>Antiplatelet use, n (%)</td>
<td>1186 (35.5)</td>
<td>1655 (32.7)</td>
<td>0.005</td>
</tr>
<tr>
<td>Warfarin use, n (%)</td>
<td>404 (6.5)</td>
<td>793 (13.0)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>Hypertension, n (%)</td>
<td>3404 (53.0)</td>
<td>4564 (75.3)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>Male sex, n (%)</td>
<td>3590 (52.6)</td>
<td>3562 (56.6)</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>Age, median (IQR)</td>
<td>66 (56-75)</td>
<td>67 (56-78)</td>
<td>&lt; 0.001</td>
</tr>
</tbody>
</table>

*χ² test and Mann Whitney u test; CV: cerebrovascular
eFigure. Distribution of Subjects in Principal Component (PA) Space and the Relation with Self-identified Race and Ethnicity (Ellipses Cluster 95% of PC Analysis Data)
REFERENCES:


9. Pezzini A, Grassi M, Iacoviello L, et al. Serum cholesterol levels, HMG-CoA reductase inhibitors and the risk of intracerebral haemorrhage. The Multicenter Study on Cerebral...


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